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STOCHASTIC ORDERING OF RANDOM k TH RECORD VALUES

Abstract. Let X_1, X_2, \dots be a sequence of independent and identically distributed random variables with continuous distribution function $F(x)$. Denote by $X(1, k), X(2, k), \dots$ the k th record values corresponding to X_1, X_2, \dots . We obtain some stochastic comparison results involving the random k th record values $X(N, k)$, where N is a positive integer-valued random variable which is independent of the X_i .

1. Introduction. Let X_1, X_2, \dots be a sequence of independent and identically distributed (i.i.d.) random variables with continuous distribution function $F(x) = P(X_1 \leq x)$. Denote by $X_{1,n} \leq \dots \leq X_{n,n}$ the order statistics of X_1, \dots, X_n . For a fixed integer $k \geq 1$, we define the corresponding k th record times, $\{L(n, k), n \geq 1\}$, and k th record values, $\{X(n, k), n \geq 1\}$, by setting

$L(1, k) = k, \quad L(n+1, k) = \min\{j > L(n, k) : X_j > X_{j-k, j-1}\} \quad \text{for } n \geq 1,$
and

$$X(n, k) = X_{L(n, k) - k + 1, L(n, k)} \quad \text{for } n \geq 1.$$

The study of k th record values and k th record times was initiated by Dziubdziela and Kopociński [2] (see e.g. Deheuvels [1] and Nevzorov [11] and the references therein). In the literature there are numerous results concerning partial ordering in the case of order statistics and record values. Interesting results for the latter are stated in Gupta and Kirmani [5], Kochar [10] and Kamps [6].

Let N be a positive integer-valued random variable which is independent of the X_i . The random variables $X(N, k)$ are called the *random k th record*

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values. The exact and limit distributions of $X(N, k)$ have been studied in the literature (see Freudentberg and Szynal [3], Grudzień [4]). The purpose of the present note is to obtain stochastic comparison results involving $X(N_1, k)$ and $X(N_2, k)$, where N_1 and N_2 are two positive integer-valued random variables which are independent of the X_i . The case when $k = 1$ was considered in Kirmani and Gupta [9]. We are concerned with the following stochastic relations: the likelihood ratio ordering and the hazard rate ordering. Convenient references for further properties and uses of these stochastic orders are Shaked and Shanthikumar [12], Shanthikumar and Yao [13], and the references therein.

2. Preliminaries. Let U and W be continuous random variables with densities $f_U(\cdot)$ and $f_W(\cdot)$, respectively. Let $F_U(\cdot)$ and $F_W(\cdot)$ be their respective distribution functions, and let $\bar{F}_U(\cdot) = 1 - F_U(\cdot)$ and $\bar{F}_W(\cdot) = 1 - F_W(\cdot)$ be their respective survival functions. Throughout, $\stackrel{d}{=}$ denotes equality in distribution; “increasing” and “decreasing” are used in the non-strict sense.

First recall the following definition and known results concerning the likelihood ratio ordering (\geq_{lr}) and the hazard rate ordering (\geq_{hr}).

We say that

$$U \geq_{lr} W \quad \text{if} \quad \frac{f_U(x)}{f_W(x)} \text{ is increasing for all } x,$$

$$U \geq_{hr} W \quad \text{if} \quad \frac{\bar{F}_U(x)}{\bar{F}_W(x)} \text{ is increasing for all } x.$$

It is well known (see for instance Shaked and Shanthikumar [12]) that

$$U \geq_{lr} W \quad \text{implies that} \quad U \geq_{hr} W.$$

The following results regarding the preservation of the order \geq_{lr} under monotone transformations is proved in Keilson and Sumita [8].

THEOREM 1. *If $U \geq_{lr} W$ and ψ is any increasing [resp. decreasing] function, then $\psi(U) \geq_{lr} [\leq_{lr}] \psi(W)$.*

Recall that a random variable U has the PF₂ (Pólya frequency of order 2) property, denoted as $U \in \text{PF}_2$, if $f_U(x)/f_U(x+a)$ is increasing in x for any given $a \geq 0$ (Karlin [7]).

The following version of the closure properties of random sums will be useful for our needs.

THEOREM 2. *Let U_1, U_2, \dots and W_1, W_2, \dots be two sequences of independent random variables such that for all i : $U_i \geq 0$ (a.s.), $U_i \geq_{lr} W_i$, and either $U_i \in \text{PF}_2$ or $W_i \in \text{PF}_2$ (or both). For $r = 1, 2, \dots$, let S_r and T_r be two independent random variables such that $S_r \stackrel{d}{=} \sum_{i=1}^r U_i$ and $T_r \stackrel{d}{=} \sum_{i=1}^r W_i$;*

and let f_{S_r} and f_{T_r} denote their respective density functions. If for all $u \geq v$ and $m \geq n$,

$$(1) \quad f_{S_m}(u)f_{T_n}(v) - f_{S_m}(v)f_{T_n}(u) \geq f_{S_n}(v)f_{T_m}(u) - f_{S_n}(u)f_{T_m}(v),$$

then

$$\sum_{i=1}^M U_i \geq_{lr} \sum_{i=1}^N W_i,$$

where M and N are any two positive integer-valued random variables which are independent of the $\{U_i\}$ and $\{W_i\}$, respectively, and $M \geq_{lr} N$.

Proof. See Shanthikumar and Yao [13].

3. Distribution of k th record values. Let X_1, X_2, \dots be i.i.d. random variables with common continuous distribution function $F(x) = P(X_1 \leq x)$. The function $R(x)$ defined by

$$R(x) = -\log(1 - F(x)), \quad -\infty \leq x \leq \infty,$$

is called the *hazard function* corresponding to the distribution function F .

Let E_1, E_2, \dots be i.i.d. exponential with mean one. Denote by $X_E(n, k)$ the k th record values for $\{E_n, n \geq 1\}$. The random variables $X(n, k)$ and $R^{\leftarrow}(X_E(n, k))$, where $R^{\leftarrow}(y) = \inf\{s : R(s) \geq y\}$ is the inverse of R , have the same distribution which we write as

$$X(n, k) \stackrel{d}{=} R^{\leftarrow}(X_E(n, k)).$$

It is known that $\{X_E(n, k), n \geq 1\}$ are the points of a homogeneous Poisson process on $(0, \infty)$ with rate k (see, for example, Deheuvels [1]). Thus

$$P(X(n, k) \leq x) = P(R^{\leftarrow}(X_E(n, k)) \leq x) = P(X_E(n, k) \leq R(x)),$$

so that

$$P(X(n, k) \leq x) = P\left(\sum_{j=1}^n Y_j \leq R(x)\right),$$

where Y_1, Y_2, \dots are i.i.d. exponential with mean k^{-1} .

Note that (Dziubdziela and Kopociński [2])

$$P(X(n, k) \leq x) = \int_0^{kR(x)} \frac{1}{(n-1)!} u^{n-1} e^{-u} du.$$

REMARK 1. It is easy to see that for $n \geq 1$ the k th record value $X(n, k)$ from an i.i.d. sequence with continuous distribution function $F(x)$ and the first record value $\tilde{X}(n, 1)$ from an i.i.d. sequence with distribution function $1 - [1 - F(x)]^k$ have the same distribution.

Let N be a positive integer-valued random variable which is independent of the X_n and let p_N denote the discrete density function of N . Then the distribution function of $X(N, k)$ is given by

$$\begin{aligned} P(X(N, k) \leq x) &= \sum_{n=1}^{\infty} P(X(n, k) \leq x) p_N(n) \\ &= \sum_{n=1}^{\infty} P\left(\sum_{j=1}^n Y_j \leq R(x)\right) p_N(n), \end{aligned}$$

where Y_1, Y_2, \dots are i.i.d. exponential with mean k^{-1} which are independent of the N . Therefore

$$P(X(N, k) \leq x) = P\left(\sum_{j=1}^N Y_j \leq R(x)\right),$$

and

$$(2) \quad X(N, k) \stackrel{d}{=} R^{\leftarrow}\left(\sum_{j=1}^N Y_j\right).$$

4. Stochastic comparisons of random k th record values. In the following theorem we obtain some comparison results involving the random k th record values.

THEOREM 3. *Let X_1, X_2, \dots be a sequence of independent and identically distributed random variables. Let N_1 and N_2 be two positive integer-valued random variables which are independent of the X_i . If $k_1 \leq k_2$ and $N_1 \geq_{lr} N_2$, then*

$$(3) \quad X(N_1, k_1) \geq_{lr} X(N_2, k_2).$$

Proof. From (2) we have

$$(4) \quad X(N_i, k_i) \stackrel{d}{=} R^{\leftarrow}\left(\sum_{j=1}^{N_i} Y_j^{(i)}\right), \quad i = 1, 2,$$

where $Y_1^{(i)}, Y_2^{(i)}, \dots$ are i.i.d. exponential with mean k_i^{-1} and are independent of the N_i . Note that $Y_j^{(i)} \in \text{PF}_2$ for $i = 1, 2$. From $k_1 \leq k_2$, it follows that $Y_j^{(1)} \geq_{lr} Y_j^{(2)}$. Writing $S_r \stackrel{d}{=} \sum_{i=1}^r Y_i^{(1)}$ and $T_r \stackrel{d}{=} \sum_{i=1}^r Y_i^{(2)}$, it is easy to see that S_r and T_r have the Erlang distributions with parameters (r, k_1^{-1}) and (r, k_2^{-1}) , respectively, and that (1) holds. Thus the conditions of Theorem 2 are satisfied, so that $\sum_{i=1}^{N_1} Y_i^{(1)} \geq_{lr} \sum_{i=1}^{N_2} Y_i^{(2)}$. Finally, by (4) and Theorem 1, we obtain (3), as desired.

REMARK 2. Kochar [10] and Kamps [6] have shown that record values and k th record values are partially ordered in the sense of likelihood ratio. These results are contained in Theorem 3.

The following result is proven in Kirmani and Gupta [9].

THEOREM 4. *Let X_1, X_2, \dots be a sequence of independent and identically distributed random variables. Let N_1 and N_2 be two positive integer-valued random variables which are independent of the X_i . If $N_1 \geq_{\text{hr}} N_2$, then*

$$(5) \quad X(N_1, 1) \geq_{\text{hr}} X(N_2, 1).$$

From Theorem 4 and Remark 2, we obtain the following particular variant of Theorem 3 for the hazard rate ordering.

COROLLARY 1. *Let X_1, X_2, \dots be a sequence of independent and identically distributed random variables. Let N_1 and N_2 be two positive integer-valued random variables which are independent of the X_i . If $N_1 \geq_{\text{hr}} N_2$, then*

$$(6) \quad X(N_1, k) \geq_{\text{hr}} X(N_2, k).$$

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