### M. MĘCZARSKI (Warszawa)

# STABILITY AND CONDITIONAL $\varGamma$ -MINIMAXITY IN BAYESIAN INFERENCE

Abstract. Two concepts of optimality corresponding to Bayesian robust analysis are considered: conditional  $\Gamma$ -minimaxity and stability. Conditions for coincidence of optimal decisions of both kinds are stated.

1. In Bayesian statistical inference arbitrariness of a unique prior distribution is a permanent question. Robust Bayesian inference deals with the problem of expressing uncertainty of the prior information and of quantitative consequences of this uncertainty. A natural measure is width (oscillation, diameter) of the range of a posterior quantity while the prior distribution  $\pi$  runs over a class  $\Gamma$  of probability distributions. If the oscillation of the posterior quantity is small, then the presence of robustness with respect to the prior inexactness can be assured.

A natural goal of research are optimal decisions under a specified loss function and a class  $\Gamma$  of prior distributions, with an idea of optimality related to the robustness problem. The concept of conditional  $\Gamma$ -minimax actions was considered in DasGupta and Studden [3] and Betrò and Ruggeri [1] and it was exhaustively substantiated therein. The idea of stability in Bayesian robust analysis was developed in Męczarski and Zieliński [5], with some additional results in Męczarski [4] and in Boratyńska and Męczarski [2]. The more formal description of the problem is given below.

Let  $(\mathcal{X}, \mathcal{F}, \{P_{\theta}\}_{\theta \in (\Theta, \mathcal{B})})$  be a statistical space,  $\Theta \subset \mathbb{R}$ . Let  $\Gamma$  be a class of probability distributions on  $(\Theta, \mathcal{B})$ , i.e. of prior distributions. It reflects the uncertainty of the prior. Let  $x \in \mathbb{R}$  be a given observation and  $a \in \mathcal{A}$ a decision (an action) about  $\theta$  based on x, with the action space  $\mathcal{A} \subset \mathbb{R}$ being a compact interval. We consider a loss function  $L(\theta, a)$  convex in a

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and the corresponding expected posterior loss (the posterior risk)  $\rho(\pi, a)$  of the action a under the prior  $\pi$  (it depends on x).

The concept of conditional  $\Gamma$ -minimaxity is as follows: construct an action  $a^* \in \mathcal{A}$  such that

$$\sup_{\pi \in \Gamma} \varrho(\pi, a^*) = \inf_{a \in \mathcal{A}} \sup_{\pi \in \Gamma} \varrho(\pi, a).$$

Such an action  $a^*$  is termed a conditional  $\Gamma$ -minimax (CGM) action.

The concept of stability is as follows: construct an action  $a^{\#} \in \mathcal{A}$  such that

$$\sup_{\pi \in \Gamma} \varrho(\pi, a^{\#}) - \inf_{\pi \in \Gamma} \varrho(\pi, a^{\#}) = \inf_{a \in \mathcal{A}} \{ \sup_{\pi \in \Gamma} \varrho(\pi, a) - \inf_{\pi \in \Gamma} \varrho(\pi, a) \},\$$

i.e. an action  $a^{\#}$  is said to be *stable* if it minimizes the oscillation of  $\varrho(\pi, a)$  on  $\Gamma$  with respect to  $a \in \mathcal{A}$ .

Solutions of some particular estimation problems show that CGM and stable actions may coincide. As seen in Boratyńska and Męczarski [2] a stable solution may lead to large losses of the posterior risk. For that reason such coincidence is desirable and also the stability of CGM actions seems to be a favourable property.

## 2. The following theorem characterizes CGM actions.

THEOREM 1 (Betrò and Ruggeri [1]). Assume that  $\rho(\pi, a)$  is a strictly convex function of a for each  $\pi \in \Gamma$ . Let  $\Pi_a = \{\pi^a : \rho(\pi^a, a) = \sup_{\pi \in \Gamma} \rho(\pi, a)\}$ be the set of least favourable priors for a decision a. Let  $a_{\pi}^B$  denote the Bayes action under  $\pi$ . If at  $\hat{a} \in \mathcal{A}$  there exist  $\pi_1$  and  $\pi_2$  in  $\Pi_{\hat{a}}$  such that  $a_{\pi_1}^B \leq \hat{a} \leq a_{\pi_2}^B$ , then  $\hat{a}$  is a CGM action.

Therefore if the stable action satisfies the conditions of Theorem 1 then it is CGM. The problem is: when is the CGM action stable?

We assume hereafter that  $\Gamma = \{\pi_{\alpha} : \alpha \in [\alpha_1, \alpha_2]\}.$ 

THEOREM 2. Let

$$(\forall a \in \mathcal{A})(\forall \pi_{\alpha} \in \Gamma) \quad \varrho(\pi_{\alpha}, a) = r(\alpha, a) = (A\alpha + B - a)^2 + C\alpha + D,$$

with  $A \neq 0, B, C, D$  real constants. If  $\Pi_{\hat{a}} = \{\pi_{\alpha_1}, \pi_{\alpha_2}\}$  for an action  $\hat{a} \in \mathcal{A}$ , then  $\hat{a}$  is stable; conversely,  $\Pi_{a^{\#}} = \{\pi_{\alpha_1}, \pi_{\alpha_2}\}$ .

Proof. Let  $\overline{\alpha} = (\alpha_1 + \alpha_2)/2$  and  $\alpha_{\min}(a)$  be the minimum point of the function  $r(\cdot, a)$ . The elementary geometry of the quadratic curve implies that the oscillation of  $r(\cdot, a)$  over  $[\alpha_1, \alpha_2]$  is the least iff  $\alpha_{\min}(a) = \overline{\alpha}$  or, equivalently,  $r(\alpha_1, a) = r(\alpha_2, a)$ , which is equivalent to  $\Pi_a = \{\pi_{\alpha_1}, \pi_{\alpha_2}\}$  because of the form of r. This yields the value  $a^{\#}$  which is the unique

solution. Since

$$\sup_{\pi \in \Gamma} \varrho(\pi, a) = \sup_{\alpha \in [\alpha_1, \alpha_2]} r(\alpha, a) = \begin{cases} r(\alpha_1, a) & \text{if } \alpha_{\min}(a) \ge \overline{\alpha}, \\ r(\alpha_2, a) & \text{if } \alpha_{\min}(a) \le \overline{\alpha}, \end{cases}$$

the conditions  $\Pi_{\hat{a}} = \{\pi_{\alpha_1}, \pi_{\alpha_2}\}$  and  $r(\alpha_1, \hat{a}) = r(\alpha_2, \hat{a})$  are equivalent, which ends the proof.

COROLLARY 1. Assume the conditions of Theorem 2. If the CGM action satisfies the conditions of Theorem 1, then it is stable. If the stable action is in the interval with endpoints  $a_{\pi_{\alpha_1}}^B$  and  $a_{\pi_{\alpha_2}}^B$ , then it is CGM.

The following examples show that the situation from Theorem 2 is realistic.

EXAMPLE 1 (Męczarski and Zieliński [5], Betrò and Ruggeri [1]). The problem is to estimate the mean  $\lambda$  in the Poisson distribution  $\mathcal{P}(\lambda)$ , given an observation x, under the prior gamma  $\mathcal{G}(\alpha, \beta)$ ,  $\alpha \in [\alpha_1, \alpha_2]$ ,  $\beta$  fixed and under the quadratic loss function. Thus

$$\begin{split} \Gamma &= \{\mathcal{G}(\alpha,\beta) : \alpha \in [\alpha_1,\alpha_2], \beta \text{ is fixed} \}\,,\\ \varrho(\pi_\alpha,a) &= r(\alpha,a) = \left(\frac{\alpha+x}{\beta+1} - a\right)^2 + \frac{\alpha+x}{(\beta+1)^2}\,,\\ a^B_{\pi_\alpha} &= \frac{\alpha+x}{\beta+1}\,. \end{split}$$

We have  $\Pi_{\hat{a}} = \{\pi_{\alpha_1}, \pi_{\alpha_2}\}$  only for  $\hat{a} = (\overline{\alpha} + 1/2 + x)/(\beta + 1)$  and this is a stable action; if  $a^B_{\pi_{\alpha_1}} \leq \hat{a} \leq a^B_{\pi_{\alpha_2}}$ , then  $\hat{a}$  is a CGM action.

EXAMPLE 2 (DasGupta and Studden [3], Boratyńska and Męczarski [2]). The problem is to estimate the mean  $\theta$  in the normal distribution  $N(\theta, b^2)$ with known b > 0, under the normal prior  $N(\mu, \sigma^2)$ , where  $\sigma \in [\sigma_1, \sigma_2] \subset \mathbb{R}_+$  and under the quadratic loss. For  $\alpha = (b^{-2} + \sigma^{-2})^{-1}$  we can write  $\Gamma = \{\pi_\alpha : \alpha \in [\alpha_1, \alpha_2]\}$ . Then

$$\varrho(\pi_{\alpha}, a) = r(\alpha, a) = \left(\alpha \frac{x - \mu}{b^2} + \mu - a\right)^2 + \alpha$$

and

$$a^B_{\pi_\alpha} = \alpha \frac{x-\mu}{b^2} + \mu \,.$$

We have  $\Pi_{\hat{a}} = \{\pi_{\alpha_1}, \pi_{\alpha_2}\}$  only for

$$\widehat{a} = \overline{\alpha} \frac{x-\mu}{b^2} + \frac{1}{2} \frac{b^2}{x-\mu}$$

and this is a stable action; if  $\hat{a}$  is in the closed interval with endpoints  $a^B_{\pi_{\alpha_1}}$ and  $a^B_{\pi_{\alpha_2}}$ , then  $\hat{a}$  is a CGM action. The last conclusions are valid for  $x \neq \mu$ ;

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otherwise any a is stable and  $a=\mu$  is CGM; therefore this CGM action is stable as well.  $\blacksquare$ 

**3.** In Theorem 2 and Examples 1 and 2 the class  $\Gamma$  is defined by a real parameter  $\alpha$  from a compact interval. The posterior risk is a quadratic function of  $\alpha$ . By minimizing the oscillation of that function we arrive at the condition  $\varrho(\pi_{\alpha_1}, a) = \varrho(\pi_{\alpha_2}, a)$ , which defines the stable action. From the shape of the function we obtain the conditions of Theorem 1 and the stability of the CGM action.

Let us consider the following example.

EXAMPLE 3 (Męczarski [4]). The problem is to estimate the parameter  $\theta = e^{-\lambda}$  in the Poisson distribution  $\mathcal{P}(\lambda)$ . The class of priors and the loss function are as in Example 1. Then

$$\varrho(\pi_{\alpha}, a) = r(\alpha, a) = (e^{\beta_1(\alpha+x)} - a)^2 - e^{2\beta_1(\alpha+x)} + e^{\beta_2(\alpha+x)},$$
  
$$\beta_i = \log \frac{\beta+1}{\beta+1+i}, \quad i = 1, 2 \text{ and } a^B_{\pi_{\alpha}} = e^{\beta_1(\alpha+x)}.$$

By looking for  $\inf_{a \in \mathcal{A}} \sup_{\pi \in \Gamma} \varrho(\pi, a)$  we obtain  $\Pi_{\hat{a}} = \{\pi_{\alpha_1}, \pi_{\alpha_2}\}$  only for the  $\hat{a}$  which is the unique solution of  $r(\alpha_1, \hat{a}) = r(\alpha_2, \hat{a})$ . If  $a_{\pi_{\alpha_2}}^B \leq \hat{a} \leq a_{\pi_{\alpha_1}}^B$ , then  $\hat{a}$  is a CGM action. By investigating the monotonicity of the oscillation of  $r(\cdot, a)$  over  $[\alpha_1, \alpha_2]$  we conclude that  $a^{\#}$  is its unique minimum point (a stable action) if and only if it satisfies  $r(\alpha_1, a^{\#}) = r(\alpha_2, a^{\#})$ .

Example 3 shows that the function  $r(\cdot, a)$  need not be quadratic to obtain the considered connection between stability and conditional  $\Gamma$ -minimaxity. As before, the condition  $r(\alpha_1, a) = r(\alpha_2, a)$  is a tool to construct an action with both properties.

THEOREM 3. Define a function r by  $r(\alpha, a) = \rho(\pi_{\alpha}, a)$ . Assume that it satisfies the following conditions:

(a)  $r(\alpha, \cdot)$  is strictly convex for any  $\alpha$ ;

(b) for any a the minimum point  $\alpha_{\min}(a)$  of  $r(\cdot, a)$  is unique and  $\alpha_{\min}$  is a strictly monotone function of a;

(c) for any  $\widetilde{\alpha}$  and  $\widetilde{a}$  such that  $\alpha_{\min}(\widetilde{a}) = \widetilde{\alpha}$  we have

$$(\forall a' < a'' \leq \widetilde{a}) \quad \frac{r(\widetilde{\alpha}, a'') - r(\widetilde{\alpha}, a')}{a'' - a'} < \frac{r(\alpha_{\min}(a''), a'') - r(\alpha_{\min}(a'), a')}{a'' - a'}$$

and

$$(\forall a'' > a' \ge \widetilde{a}) \quad \frac{r(\widetilde{\alpha}, a'') - r(\widetilde{\alpha}, a')}{a'' - a'} > \frac{r(\alpha_{\min}(a''), a'') - r(\alpha_{\min}(a'), a')}{a'' - a'};$$

(d) the function  $r(\alpha_1, a) - r(\alpha_2, a)$  is monotone in a.

Then the conclusion of Theorem 2 holds.

Proof. Denote the oscillation of  $r(\cdot, a)$  over  $[\alpha_1, \alpha_2]$  by Q(a). Then

$$Q(a) = \begin{cases} r(\alpha_{2}, a) - r(\alpha_{1}, a) & \text{if } \alpha_{\min}(a) \leq \alpha_{1}, \\ r(\alpha_{2}, a) - r(\alpha_{\min}(a), a) & \text{if } \alpha_{1} < \alpha_{\min}(a) < \alpha_{2} \\ & \text{and } r(\alpha_{1}, a) - r(\alpha_{2}, a) \leq 0, \\ r(\alpha_{1}, a) - r(\alpha_{\min}(a), a) & \text{if } \alpha_{1} < \alpha_{\min}(a) < \alpha_{2} \\ & \text{and } r(\alpha_{1}, a) - r(\alpha_{2}, a) \geq 0, \\ r(\alpha_{1}, a) - r(\alpha_{2}, a) & \text{if } \alpha_{\min}(a) \geq \alpha_{2}. \end{cases}$$

Let  $a_i$  be a solution of  $\alpha_{\min}(a) = \alpha_i$ , i = 1, 2. The type of monotonicity of the function  $r(\alpha_1, a) - r(\alpha_2, a)$  agrees with that of  $\alpha_{\min}$ . Then for  $\alpha_{\min}$ increasing we have

$$Q(a) = \begin{cases} r(\alpha_2, a) - r(\alpha_1, a) & \text{if } a \le a_1, \\ r(\alpha_2, a) - r(\alpha_{\min}(a), a) & \text{if } a_1 < a < a_2 \\ & \text{and } r(\alpha_1, a) - r(\alpha_2, a) \le 0, \\ r(\alpha_1, a) - r(\alpha_{\min}(a), a) & \text{if } a_1 < a < a_2 \\ & \text{and } r(\alpha_1, a) - r(\alpha_2, a) \ge 0, \\ r(\alpha_1, a) - r(\alpha_2, a) & \text{if } a \ge a_2 \end{cases}$$

and for  $\alpha_{\min}$  decreasing we have

$$Q(a) = \begin{cases} r(\alpha_1, a) - r(\alpha_2, a) & \text{if } a \le a_2, \\ r(\alpha_1, a) - r(\alpha_{\min}(a), a) & \text{if } a_2 < a < a_1 \\ & \text{and } r(\alpha_1, a) - r(\alpha_2, a) \ge 0, \\ r(\alpha_2, a) - r(\alpha_{\min}(a), a) & \text{if } a_2 < a < a_1 \\ & \text{and } r(\alpha_1, a) - r(\alpha_2, a) \le 0, \\ r(\alpha_2, a) - r(\alpha_1, a) & \text{if } a \ge a_1. \end{cases}$$

Observe that the functions  $r(\alpha_i, a) - r(\alpha_{\min}(a), a)$  have their unique minimum points at  $a_i$  and they are decreasing for  $a < a_i$  and increasing for  $a > a_i, i = 1, 2$ , respectively. This implies that Q has its unique minimum point at  $a^{\#}$  which is defined by the condition  $r(\alpha_1, a^{\#}) - r(\alpha_2, a^{\#}) = 0$ .

Now,

$$\sup_{\alpha \in [\alpha_1, \alpha_2]} r(\alpha, a) = \begin{cases} r(\alpha_1, a) & \text{if } r(\alpha_1, a) \ge r(\alpha_2, a), \\ r(\alpha_2, a) & \text{if } r(\alpha_1, a) \le r(\alpha_2, a) \end{cases}$$

and therefore  $\Pi_{a^{\#}} = \{\pi_{\alpha_1}, \pi_{\alpha_2}\}$ . Conversely, if  $\Pi_{\hat{a}} = \{\pi_{\alpha_1}, \pi_{\alpha_2}\}$ , then  $\widehat{a} = a^{\#}$ . This ends the proof.

COROLLARY 2. Assume the conditions of Theorem 3. Let the function  $f(\alpha) = a^B_{\pi_{\alpha}}$  be strictly monotone and  $f([\alpha_1, \alpha_2]) = \mathcal{A}$ . Then the conclusion of Corollary 1 holds.

Now we can see that the conditions of Theorem 3 and Corollary 2 hold in Examples 1, 2 and 3.

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MAREK MĘCZARSKI INSTITUTE OF ECONOMETRICS WARSAW SCHOOL OF ECONOMICS AL. NIEPODLEGŁOŚCI 162 02-554 WARSZAWA, POLAND

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