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TRANSITION PROBABILITY DENSITY FUNCTION  
OF A CERTAIN DIFFUSION PROCESS  
CONCENTRATED ON A HALF LINE

*Abstract.* We prove that under some assumptions there exists a diffusion process satisfying a one-dimensional Itô equation and living in a time-dependent half line. We give a formula on the transition probability density function of this process. This is also a probabilistic formula for a solution of a deterministic Fokker-Planck equation in a time-dependent half line.

**1. Introduction.** Diffusion processes are used in biology to the description of population dynamics. In [5] and [6] A. G. Nobile and L. M. Ricciardi consider a growth process of a population in a fluctuating environment. This process can be a solution of some one-dimensional Itô or Stratonovich equation. In some cases the transition probability density of these solutions, concentrated on a finite interval or on a half line, are given ([5]). Some sufficient conditions on the existence of a strong solution of the Itô equation

$$(1.1) \quad X(t) = X_0 + \int_0^t a(s, X(s)) ds + \int_0^t b(s, X(s)) dW(s),$$

living in a bounded region in  $\mathbb{R}^n$ , are given in [3]. The transition probability density functions for one-dimensional diffusion processes with varying boundaries are investigated in [2]. For a one-dimensional process satisfying (1.1), the transition probability density function concentrated on a finite spatial interval is given in [4].

In the present paper we give a formula for the transition probability density function of a solution of (1.1), concentrated on a time-dependent half

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line. Using our result we obtain the transition probability density function for the growth process satisfying the equation

$$(1.2) \quad dX(t) = \left[ -\rho X \left(1 - \frac{\ln X}{\beta}\right) + \lambda X \ln X \left(1 - \frac{\ln X}{\beta}\right) \right] dt \\ - X \left(1 - \frac{\ln X}{\beta}\right) dW(t)$$

in the diffusion interval  $(\exp(\beta), \infty)$ , in the case  $\beta > 0$  and  $\lambda = -1/(2\beta)$ . This equation is investigated in [6], p. 293, but the transition probability density functions are not given there.

**2. Constant boundary.** To formulate our theorem concerning a diffusion process concentrated on a half line we need some definitions.

Let the coefficients  $a(t, x)$  and  $b(t, x)$  of (1.1) be defined for  $t \geq 0$  and  $x \geq 0$ . If  $b(t, x) \neq 0$  for  $t \geq 0$  and  $x > 0$ , then we introduce transformations  $u(t, x)$  and  $v(t, x)$  by

$$(2.1) \quad u(t, x) = \int_1^x \frac{dy}{b(t, y)},$$

$$(2.2) \quad x = \int_1^{v(t, x)} \frac{dy}{b(t, y)}.$$

Note that  $v(t, \cdot)$  is the inverse to  $u(t, \cdot)$  for  $t \geq 0$ . Then we define

$$(2.3) \quad m(t, x) = u_t(t, v(t, x)) + u_x(t, v(t, x))a(t, v(t, x)) \\ + \frac{1}{2}u_{xx}(t, v(t, x))b^2(t, v(t, x))$$

$$(2.4) \quad \sigma(t, x) = u_x(t, v(t, x))b(t, v(t, x)).$$

**THEOREM 1.** Assume that the random variable  $X_0$  is independent of a given Wiener process  $W(t)$  and  $X_0 > 0$  with probability 1. Let the coefficients  $a(t, x)$  and  $b(t, x)$  satisfy the following conditions:

1)  $a(t, x)$  and  $b(t, x)$  are  $C^2$  in some open neighbourhood of the set  $\{(t, x) : t \geq 0, x \geq 0\}$ .

2)  $b(t, 0) = 0$  and  $b_x(t, 0) \neq 0$  for  $t \geq 0$ ,  $b(t, x) > 0$  for  $t \geq 0, x > 0$ , and

$$\int_1^\infty \frac{ds}{b(t, s)} = \infty \quad \text{for each } t \geq 0.$$

3)  $m_x(t, x)$  and  $m_t(t, x)$  are bounded for  $t \geq 0, x \in \mathbb{R}$ .

4)  $\mathbb{E}[u(0, X_0)]^2 < \infty$ .

Then (1.1) has a strong solution  $X(t)$  satisfying the initial condition  $X(t)|_{t=0} = X_0$  a.s. and such that  $\text{Prob}\{X(t) > 0, t \geq 0\} = 1$ . If  $X(t)$  and

$Y(t)$  are two such solutions of (1.1), then

$$\text{Prob}\{\sup_{t \geq 0} |X(t) - Y(t)| = 0\} = 1,$$

i.e. pathwise uniqueness holds.

A transition probability density of the process  $X(t)$  exists and is given by

$$(2.5) \quad q(t, x, s, y) = \frac{1}{\sqrt{2\pi(s-t)b(s, y)}} \\ \times \exp \left\{ -\frac{\{u(s, y) - u(t, x)\}^2}{2(s-t)} + \bar{M}(s, u(s, y)) - \bar{M}(t, u(t, x)) \right\} \\ \times \mathbb{E} \exp \left\{ (s-t) \int_0^1 \bar{B}(t+z(s-t), u(t, x)) \right. \\ \left. + \sqrt{s-t} \eta_{t,s}^*(z) + z\{u(s, y) - u(t, x)\} dz \right\},$$

where

$$(2.6) \quad \bar{B}(t, x) = -\frac{1}{2}m^2(t, x) - \frac{1}{2}m'_x(t, x) - \int_0^x m'_t(t, y) dy,$$

$$(2.7) \quad \bar{M}(t, u(t, x)) = \int_1^x \left( \int_1^y \frac{\partial}{\partial t} \left( \frac{1}{b(t, s)} \right) ds + \frac{a(t, y)}{b(t, y)} - \frac{1}{2}b_x(t, y) \right) \frac{dy}{b(t, y)},$$

$$W_{t,s}^* = \frac{W[t+(s-t)z] - W(t)}{\sqrt{s-t}}, \quad \eta_{t,s}^*(z) = W_{t,s}^*(z) - zW_{t,s}^*(1),$$

$u(t, x)$  is given by (2.1) and  $m(t, x)$  is given by (2.3).

*Proof.* We consider the function  $u(t, \cdot)$  for  $t \geq 0$ . By Conditions 1 and 2 we have  $b(t, y) = b(t, y) - b(t, 0) = yb_x(t, \xi)$  for some  $\xi \in (0, y)$  and  $0 < b_x(t, \xi) \leq M < \infty$  for some  $M \in \mathbb{R}$ . Hence

$$\frac{1}{b(t, y)} = \frac{1}{yb_x(t, \xi)} \geq \frac{1}{yM}.$$

Thus

$$\int_x^1 \frac{dy}{b(t, y)} \geq \frac{1}{M} \int_x^1 \frac{dy}{y} = \frac{1}{M} \ln \frac{1}{x},$$

$$\int_x^1 \frac{dy}{b(t, y)} \rightarrow +\infty \quad \text{and} \quad u(t, x) \rightarrow -\infty \quad \text{as} \quad x \rightarrow 0^+.$$

By Condition 2,  $u(t, x) \rightarrow \infty$  as  $x \rightarrow \infty$ , hence  $u(t, x)$  is a one-to-one mapping from  $(0, \infty)$  onto  $\mathbb{R}$ , for  $t \geq 0$ .

Now we look for a process  $\xi(t)$  satisfying (1.1) with the coefficients  $m(t, x)$  and  $\sigma(t, x)$  given by (2.3) and (2.4) and with the initial condition  $P\{\xi(t)|_{t=0} = u(0, X_0)\} = 1$ .

To this end we show that  $m(t, x)$  and  $\sigma(t, x)$  satisfy all the assumptions of the existence and uniqueness theorem ([1], Theorem 1, p. 40). Differentiating the identity (2.2) with respect to  $x$  we obtain

$$(2.8) \quad v_x(t, x) = b(t, v(t, x)).$$

In the same way, by the identity  $u(t, v(t, x)) = x$ , we obtain  $u_x(t, v(t, x)) \times v_x(t, x) = 1$ , and consequently

$$(2.9) \quad u_x(t, v(t, x)) = \frac{1}{b(t, v(t, x))}.$$

Thus, from (2.4) and (2.9), we have

$$(2.10) \quad \sigma(t, x) = 1.$$

By (2.10) and by Conditions 3 and 4 we conclude that all assumptions of ([1], Theorem 1, p. 40) are satisfied. Thus there exists a solution  $\xi(t)$  of (1.1) satisfying the conditions

(A)  $\xi(t)$  is a.s. continuous and  $\xi(t) = u(0, X_0)$  for  $t = 0$ .

(B)  $\sup \mathbb{E}(\xi(t))^2 < \infty$ .

If  $\xi_1(t)$  and  $\xi_2(t)$  are two solutions of (1.1) satisfying (A) and (B), then

$$\text{Prob}\left\{\sup_{0 \leq t \leq T} |\xi_1(t) - \xi_2(t)| = 0\right\} = 1,$$

i.e. pathwise uniqueness holds.

By Condition 1 and by (2.1), the coefficients  $m(t, x)$  and  $\sigma(t, x)$  are continuous in both arguments. Hence by ([1], Theorem 2, p. 68) the process  $\xi(t)$  is a diffusion with diffusion coefficient  $\sigma^2(t, x)$  and displacement coefficient  $m(t, x)$ .

Now we show that all conditions of ([1], Theorem 2, p. 96) are satisfied. By (2.1), (2.3), (2.10) and by Condition 3 we conclude that  $m_x(t, x)$  and  $\sigma_x(t, x)$  exist and are bounded. The derivatives  $\sigma''_{xx}$ ,  $\sigma''_{xt}$ ,  $\sigma'_t$  and  $\sigma''_{tt}$  exist. By (2.10) we can calculate the following functions (cf. the statement of ([1], Theorem 2, p. 96)):

$$(2.11) \quad x = \int_0^{h(t,x)} \frac{dy}{\sigma(t, y)} = h(t, x)$$

and

$$(2.12) \quad \bar{m}(t, x) = \int_0^{h(t,x)} \frac{\sigma_t(t, y)}{\sigma^2(t, y)} dy + \frac{m(t, h(t, x))}{\sigma(t, h(t, x))} - \frac{1}{2} \sigma_x(t, h(t, x)) = m(t, x)$$

so  $\bar{B}(t, x)$  is given by (2.6).

Taking into account Condition 3 we conclude that there exists some constant  $K > 0$  such that

$$\left| -\frac{1}{2}m'_x(t, x) - \int_0^x m'_t(t, y) dy \right| \leq K(1 + |x|).$$

Hence

$$\limsup_{|x| \rightarrow \infty} \frac{1}{1+x^2} \sup_{0 \leq t \leq T} \bar{B}(t, x) = \limsup_{|x| \rightarrow \infty} \frac{1}{1+x^2} \sup_{0 \leq t \leq T} \left( -\frac{1}{2}m^2(t, x) \right) \leq 0.$$

Thus all conditions of ([1], Theorem 2, p. 96) are satisfied.

Hence the transition probability of the process  $\xi(t)$ :

$$P(t, x, s, y) = \text{Prob}\{\xi(s) \leq y \mid \xi(t) = x\}, \quad t < s,$$

has a density given by

$$\begin{aligned} (2.13) \quad p(t, x, s, y) &= \frac{1}{\sqrt{2\pi(s-t)}\sigma(s, y)} \\ &\times \exp \left\{ -\frac{(\varphi(s, y) - \varphi(t, x))^2}{2(s-t)} + \bar{M}(s, \varphi(s, y)) - \bar{M}(t, \varphi(t, x)) \right\} \\ &\times \mathbb{E} \exp \left\{ (s-t) \int_0^1 \bar{B}(t + z(s-t), \varphi(t, x)) \right. \\ &\quad \left. + \sqrt{s-t} \eta_{t,s}^*(z) + z[\varphi(s, y) - \varphi(t, x)] \right\} dz, \end{aligned}$$

where  $h(t, x)$ ,  $\bar{m}(t, x)$  and  $\bar{B}(t, x)$  are given by (2.11), (2.12) and (2.6) respectively, and moreover

$$\begin{aligned} \bar{M}(t, x) &= \int_0^x \bar{m}(t, y) dy, \quad \varphi(t, x) = \int_0^x \frac{dy}{\sigma(t, y)}, \\ W_{t,s}^* &= \frac{W(t + (s-t)z) - W(t)}{\sqrt{s-t}}, \quad \eta_{t,s}^*(z) = W_{t,s}^*(z) - zW_{t,s}^*(1). \end{aligned}$$

Now we define the process

$$X(t) = v(t, \xi(t)).$$

Using Itô's formula ([1], Theorem 4, p. 24) it is easy to prove that the process  $X(t)$  satisfies the equation (1.1) with coefficients  $a(t, x)$  and  $b(t, x)$ . Thus  $X(t)$  is a strong solution of (1.1), and moreover

$$X(0) = v(0, u(0, X_0)) = X_0.$$

If  $Y(t)$  is a solution of (1.1) and  $Y(0) = X_0$ , then

$$\begin{aligned} \text{Prob}\left\{\sup_{0 \leq t \leq T} |X(t) - Y(t)| = 0\right\} \\ = \text{Prob}\left\{\sup_{0 \leq t \leq T} |u(t, X(t)) - u(t, Y(t))| = 0\right\} = 1, \end{aligned}$$

i.e. pathwise uniqueness holds.

Furthermore,

$$\text{Prob}\{X(t) > 0, t \geq 0\} = 1.$$

The process  $X(t)$  is a diffusion with diffusion coefficient  $b^2(t, x)$  and drift coefficient  $a(t, x)$  ([1], p. 66). By the definition of the process  $X(t)$ , and by Lemma 1 of [4], the transition probability

$$Q(t, x, s, y) = \text{Prob}\{X(s) \leq y \mid X(t) = x\}, \quad t < s,$$

has a density given by

$$(2.14) \quad q(t, x, s, y) = p(t, u(t, x), s, u(s, y))u_y(s, y)$$

where  $p(t, x, s, y)$  is given by (2.13).

Now we want to express the density (2.14) in terms of the coefficients of (1.1). To this end we make some calculations. The functions  $\varphi(t, x)$  and  $\bar{M}(t, x)$  are as follows:

$$\begin{aligned} \varphi(t, x) &= \int_0^x \frac{dy}{\sigma(t, y)} = \int_0^x dy = x, \\ \bar{M}(t, x) &= \int_0^x \bar{m}(t, y) dy = \int_0^x m(t, y) dy. \end{aligned}$$

After some standard calculations we find that

$$(2.15) \quad m(t, x) = \int_1^{v(t, x)} \frac{\partial}{\partial t} \left( \frac{1}{b(t, s)} \right) ds + \frac{a(t, v(t, x))}{b(t, v(t, x))} - \frac{1}{2} b_x(t, v(t, x)).$$

Next we calculate

$$\begin{aligned} \bar{M}(t, u(t, x)) &= \int_0^{u(t, x)} \left( \int_1^{v(t, z)} \frac{\partial}{\partial t} \left( \frac{1}{b(t, s)} \right) ds \right. \\ &\quad \left. + \frac{a(t, v(t, z))}{b(t, v(t, z))} - \frac{1}{2} b_x(t, v(t, z)) \right) dz. \end{aligned}$$

We substitute in this integral  $v(t, z) = y$  (for fixed  $t$ ). From the identity  $v(t, u(t, z)) = z$ , we obtain  $z = u(t, y)$ . By (2.1) we have  $u_x(t, x) = 1/b(t, x)$ , hence  $dz = dy/b(t, y)$ , and consequently we obtain the formula (2.7) for  $\bar{M}(t, u(t, x))$ .

From (2.6) we find that  $\bar{B}(t, x) = W(t, v(t, x))$ , where

$$\begin{aligned} W(t, z) = & -\frac{1}{2} \left( \int_1^z \frac{\partial}{\partial t} \left( \frac{1}{b(t, s)} \right) ds + \frac{a(t, z)}{b(t, z)} - \frac{1}{2} b_x(t, z) \right)^2 \\ & - \frac{1}{2} \left( a_x(t, z) - \frac{b_t(t, z) + a(t, z)b_x(t, z)}{b(t, z)} - \frac{1}{2} b_{xx}(t, z)b(t, z) \right) \\ & - \int_1^z \left\{ \int_1^y \frac{\partial^2}{\partial t^2} \left( \frac{1}{b(t, s)} \right) ds + \left( \frac{b_t(t, y)}{b(t, y)} - a_x(t, y) + \frac{a(t, y)b_x(t, y)}{b(t, y)} \right. \right. \\ & \left. \left. + \frac{1}{2} b_{xx}(t, y)b(t, y) \right) \int_1^y \frac{\partial}{\partial t} \left( \frac{1}{b(t, s)} \right) ds \right. \\ & \left. + \frac{a_t(t, y)}{b(t, y)} - \frac{a(t, y)b_t(t, y)}{b^2(t, y)} - \frac{1}{2} b_{tx}(t, y) \right\} \frac{dy}{b(t, y)}, \end{aligned}$$

which leads to the formula (2.5) for the transition probability density of the process  $X$ .

This completes the proof of Theorem 1.

In this way we have obtained the probabilistic formula (2.5) for a solution of the following deterministic Fokker-Planck equation:

$$\frac{\partial q(t, x, s, y)}{\partial s} = -\frac{\partial}{\partial y} [a(s, y)q(t, x, s, y)] + \frac{1}{2} \frac{\partial^2}{\partial y^2} [b^2(s, y)q(t, x, s, y)], \quad t < s,$$

with the initial condition  $q(t, x, t, y) = \delta(x - y)$ ,  $x \geq 0$ .

This solution has the following properties:  $q(t, x, s, y) \geq 0$  for  $s \geq t$  and  $y \geq 0$ ,  $q(t, x, s, y) = 0$  for  $s \geq t$  and  $y < 0$ , and moreover  $\int_0^\infty q(t, x, s, y) dy = 1$  for each  $s \geq 0$ .

**3. Time-dependent half line.** Let  $G = \{(t, x) : 0 \leq t \leq T, x \geq \gamma(t)\}$ , where  $\gamma$  is defined for  $t \in [0, T]$ . Define  $f(t, x) = x - \gamma(t)$ . So  $f(t, \cdot)$  is a one-to-one mapping from  $(\gamma(0), \infty)$  onto  $(0, \infty)$ , for  $t \in [0, T]$ . Let  $g(t, \cdot)$  denote the inverse of  $f(t, \cdot)$ , i.e.  $g(t, x) = x + \gamma(t)$ .

Assume that the random variable  $X_{1,0}$  is independent of a Wiener process  $W(t)$  and  $\text{Prob}\{X_{1,0} > \gamma(0)\} = 1$ .

Now we consider (1.1) with the coefficients  $a_1(t, x)$  and  $b_1(t, x)$  defined in  $G$ , with the initial condition  $X_1(t)|_{t=0} = X_{1,0}$  a.s. We define

$$(3.1) \quad a(t, x) = a_1(t, g(t, x)),$$

$$(3.2) \quad b(t, x) = b_1(t, g(t, x)) \quad \text{for } t \in [0, T] \text{ and } x > 0.$$

If  $a(t, x)$  and  $b(t, x)$  given by (3.1) and (3.2) satisfy all the conditions of Theorem 1, then the following corollary is true:

COROLLARY 1. Equation (1.1) with coefficients  $a_1(t, x)$  and  $b_1(t, x)$  has a strong solution  $X_1(t)$  satisfying the initial condition  $\text{Prob}\{X_1(t)|_{t=0} = X_{1,0}\} = 1$ . Moreover,  $\text{Prob}\{X_1(t) > \gamma(t), 0 \leq t \leq T\} = 1$ .

If  $X_1$  and  $Y_1$  are two such solutions of (1.1) then

$$\text{Prob}\left\{\sup_{0 \leq t \leq T} |X_1(t) - Y_1(t)| = 0\right\} = 1,$$

i.e. pathwise uniqueness holds.

A transition probability density of the process  $X_1(t)$  exists and is given by

$$(3.3) \quad q_1(t, x, s, y) = q(t, f(t, x), s, f(s, y))$$

where  $q(t, x, s, y)$  is given by (2.5) for the coefficients  $a(t, x)$  and  $b(t, x)$  defined by (3.1) and (3.2). The function  $q_1$  is a solution of the Fokker-Planck equation and satisfies the condition

$$\int_{\gamma(t)}^{\infty} q_1(t, x, s, y) dy = 1 \quad \text{for } s \geq t.$$

**4. Example.** We can use Corollary 1 to calculate the transition probability density function of the diffusion process  $X_1(t)$  satisfying the equation (1.2), in the case  $\lambda = -1/(2\beta)$ ,  $\beta > 0$ , on the interval  $(e^\beta, \infty)$ , with the initial condition  $\text{Prob}\{X_1(0) = x_0\} = 1$ , where  $x_0 > e^\beta$ . In this case  $\gamma(t) = e^\beta$  for  $t \geq 0$ . Hence

$$\begin{aligned} a(x) &= -\varrho(x + e^\beta) \left(1 - \frac{\ln(x + e^\beta)}{\beta}\right) \\ &\quad + \lambda(x + e^\beta) \ln(x + e^\beta) \left(1 - \frac{\ln(x + e^\beta)}{\beta}\right), \\ b(x) &= (x + e^\beta) \left(\frac{\ln(x + e^\beta)}{\beta} - 1\right). \end{aligned}$$

Now we can show that  $a(x)$  and  $b(x)$  satisfy all the assumptions of Theorem 1 in the interval  $(0, \infty)$ :

$$\begin{aligned} b(0) &= 0, \quad b'(0) = 1/\beta > 0, \quad b(x) > 0 \quad \text{for } x > 0, \\ \int_1^\infty \frac{dx}{b(x)} &\geq \beta \int_1^\infty \frac{dx}{(x + e^\beta) \ln(x + e^\beta)} = \infty. \end{aligned}$$

From (2.15) we have

$$(4.1) \quad m(x) = \frac{a(v(x))}{b(v(x))} - \frac{1}{2} b'(v(x)).$$



Consequently, taking into account (2.8), we obtain

$$m'(x) = \lambda + \frac{1}{2\beta} - \left( \lambda + \frac{1}{2\beta} \right) \frac{\ln(x + e^\beta)}{\beta}.$$

So  $m'(x)$  is bounded if and only if  $\lambda + 1/(2\beta) = 0$ . Then  $m'(x) \equiv 0$  and all the assumptions of Theorem 1 are satisfied. From (2.2) and (4.1) we have

$$m(x) \equiv m(0) = \varrho + \frac{1}{2} - \frac{1}{2\beta}.$$

Define

$$\mu = \varrho + \frac{1}{2} - \frac{1}{2\beta}.$$

From (2.6) and (2.12) we have, respectively,

$$(4.2) \quad \bar{B}(x) = -\frac{1}{2}\mu^2,$$

$$(4.3) \quad \bar{M}(x) = \mu x.$$

By Corollary 1 we conclude that the probability

$$\text{Prob}\{X_1(t) \leq y \mid X_1(0) = x_0\}$$

has a density function given by

$$q_1(0, x_0, s, y) = q(0, x_0 - e^\beta, s, y - e^\beta)$$

where  $q$  is given by (2.5). Taking into account (4.2) and (4.3) we have

$$q_1(0, x_0, s, y) = \frac{1}{\sqrt{2\pi s b(y - e^\beta)}} \exp \left\{ -\frac{[u(y - e^\beta) - u(x_0 - e^\beta)]^2}{2s} + \mu[u(y - e^\beta) - u(x_0 - e^\beta)] \right\} \exp(-\frac{1}{2}\mu^2 s).$$

By (2.1),

$$\begin{aligned} u(y - e^\beta) - u(x_0 - e^\beta) &= \int_{z_1}^{z_2} \frac{dz}{b(z)} = \beta \int_{z_1}^{z_2} \frac{dz}{(z + e^\beta)[\ln(z + e^\beta) - \beta]} \\ &= \beta \ln \left( \frac{\ln y - \beta}{\ln x_0 - \beta} \right), \end{aligned}$$

where  $z_1 = x_0 - e^\beta$ ,  $z_2 = y - e^\beta$ . So

$$q_1(0, x_0, s, y) = \frac{\beta \exp(-\frac{1}{2}\mu^2 s)}{\sqrt{2\pi s y (\ln x_0 - \beta)}} (q(y))^{w(y)}$$

where

$$q(y) = \frac{\ln y - \beta}{\ln x_0 - \beta}, \quad w(y) = \mu\beta - 1 - \frac{\beta^2}{2\varrho} \ln q(y).$$

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