

*Poisson-normal type* if each finite dimensional marginal probability function of  $P$  is of the Poisson-normal type.

Let us now consider the multinomial distribution given by the formula

$$(1) \quad P_n(x_1 = k_1, x_2 = k_2, \dots, x_r = k_r) = \frac{n!}{k_1! k_2! \dots k_r!} p_{n1}^{k_1} p_{n2}^{k_2} \dots p_{nr}^{k_r}$$

where  $0 < p_{nm} < 1$ ,  $p_{nm}$  are arbitrary functions of  $n$  and the  $k_m$  ( $m=1, 2, \dots, r$ ) are non-negative integers satisfying the equality

$$\sum_{m=1}^r k_m = n.$$

The variable  $(x_1, x_2, \dots, x_r)$  can be reduced with probability 1 — in view of the last equality — to an  $(r-1)$ -dimensional variable.

The following theorem has been proved by the author [1]:

**THEOREM 1.** *Let the random variable  $(x_1, x_2, \dots, x_r)$  be distributed according to (1) and let the sequence  $G_n$  of probability functions of the random variables*

$$(A_{n1}x_1 + B_{n1}, A_{n2}x_2 + B_{n2}, \dots, A_{nr}x_r + B_{nr}),$$

where  $A_{nm} \neq 0$  and  $B_{nm}$  ( $n=1, 2, 3, \dots; m=1, 2, \dots, r$ ) are real numbers, converge, as  $n \rightarrow \infty$ , to a non-singular  $(r-1)$ -dimensional probability function  $G$ . Then  $G$  is necessarily of the Poisson normal type.

Let us now modify the multinomial distribution given by (1). Namely, let us suppose that the number  $r$  in (1) — which we shall denote by  $r_n$  — is a non-decreasing function of  $n$ , increasing to infinity, as  $n \rightarrow \infty$ . The following theorem answers a question put to the author by G. Hajos:

**THEOREM 1a.** *Let the sequence  $G_n$  of probability functions of the random variables*

$$(A_{n1}x_1 + B_{n1}, A_{n2}x_2 + B_{n2}, \dots, A_{nr_n}x_{r_n} + B_{nr_n}),$$

where  $A_{nm} \neq 0$  and  $B_{nm}$  ( $n=1, 2, 3, \dots; m=1, 2, \dots, r_n$ ) are real constants, converge, as  $n \rightarrow \infty$  and  $r_n \rightarrow \infty$ , to a non-singular denumerably dimensional probability function  $G$ . Then  $G$  is necessarily of the Poisson-normal type.

**Proof.** Let the assumptions of theorem 1a be satisfied. Thus the sequences of arbitrary  $s$ -dimensional ( $s=1, 2, 3, \dots$ ) marginal probability functions of  $G_n$  converge, as  $n \rightarrow \infty$ , to the corresponding marginal probability function of the limiting probability function  $G$ . However, from theorem 1 it follows that arbitrary marginal probability functions of  $G$  are of the Poisson-normal type. Thus, taking into account definition 3a, we obtain the assertion of theorem 1a.

## A limit theorem for a modified Bernoulli scheme

by

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1. An essential role in the considerations of this paper is played by the known theorem of Kolmogorov [2], stating that the probability function  $P$  in the space  $(x_1, x_2, x_3, \dots)$  is uniquely determined by the set of all finite dimensional marginal probability functions  $P_{i_1, i_2, \dots, i_s}$  of  $P$  defined for all Borel sets in the corresponding  $s$ -dimensional space  $(x_{i_1}, x_{i_2}, \dots, x_{i_s})$  for  $s=1, 2, 3, \dots$  and for arbitrary integers  $i_1, i_2, \dots, i_s$ .

**DEFINITION 1.** We shall say that the denumerably dimensional probability function  $P$ , given in the space  $(x_1, x_2, x_3, \dots)$ , is not singular if for  $s=1, 2, 3, \dots$  and for arbitrary integers  $i_1, i_2, \dots, i_s$  the marginal probability function  $P_{i_1, i_2, \dots, i_s}$  is not singular<sup>1)</sup> in the usual sense.

**DEFINITION 2.** We shall say that the sequence  $P_n$  of denumerably dimensional probability functions in the space  $(x_1, x_2, x_3, \dots)$  converges, as  $n \rightarrow \infty$ , to a probability function  $P$  if for  $s=1, 2, \dots$  and arbitrary integers  $i_1, i_2, \dots, i_s$  the sequence  $P_{n(i_1, i_2, \dots, i_s)}$  of  $s$ -dimensional marginal probability functions of  $P_n$  converges, as  $n \rightarrow \infty$ , to the corresponding marginal probability function  $P_{i_1, i_2, \dots, i_s}$  of the limiting probability function  $P$ .

**DEFINITION 3.** The non-singular probability function  $P$  in the space  $(x_1, x_2, \dots, x_s)$  is of the *Poisson-normal type* if it is a probability function of a variable  $(\xi, \eta)$  where  $\xi$  is a  $j$ -dimensional Poisson<sup>2)</sup> variable,  $\eta$  an  $(s-j)$ -dimensional normal variable ( $0 \leq j \leq s$ ),  $\xi$  and  $\eta$  being independent.

**DEFINITION 3a.** We shall say that the non-singular denumerably dimensional probability function  $P$  in the space  $(x_1, x_2, x_3, \dots)$  is of the

<sup>1)</sup> The probability function  $P$ , defined in the space  $(x_1, x_2, \dots, x_s)$ , is called *singular* if the whole mass of probability lies in a  $z$ -dimensional hyperplane, where  $z < s$ .

<sup>2)</sup> A  $j$ -dimensional random variable  $(y_1, y_2, \dots, y_j)$  is called a *Poisson variable* if its probability function is given by the formula

$$P(y_1 = A_1 k_1 + B_1, y_2 = A_2 k_2 + B_2, \dots, y_j = A_j k_j + B_j) = \prod_{m=1}^j e^{-\lambda_m} \lambda_m^{k_m} / k_m!$$

where  $k_m = 0, 1, 2, \dots$  and  $\lambda_m > 0$ ,  $A_m \neq 0$  and  $B_m$  are real constants.

Let us now observe that, the assumption of theorem 1a being satisfied, one can assume<sup>3)</sup> that the set  $M = \{1, 2, 3, \dots\}$  of indices  $m$  can be divided into two subsets  $M_1 = \{m_1, m_2, \dots\}$  and  $M_2 = \{m_2, m_2, \dots\}$  in such a way that for  $m \in M_1$  the relation

$$(2) \quad \lim_{n \rightarrow \infty} n p_{nm} = \lambda_m,$$

where  $0 < \lambda_m < \infty$ , holds, and for  $m \in M_2$  the relations

$$(3) \quad \lim_{n \rightarrow \infty} p_{nm} = p_m, \quad \lim_{n \rightarrow \infty} n p_{nm} = \infty$$

hold. Let us introduce new variables

$$(4) \quad \begin{aligned} y_m &= x_m & (m \in M_1), \\ y_m &= (x_m - n p_{nm}) / \sqrt{n p_{nm}} & (m \in M_2). \end{aligned}$$

Let

$$\varphi_{ns} = \varphi_{ns}(t_{m_1}, \dots, t_{m_j}, t_{m_{j+1}}, \dots, t_{m_s})$$

denote the characteristic function of the random variable  $(y_{m_1}, \dots, y_{m_j}, y_{m_{j+1}}, \dots, y_{m_s})$ , where  $s$  is an arbitrary constant integer and  $m_v \in M_1$  for  $v=1, 2, 3, \dots, j$  and  $m_v \in M_2$  for  $v=j+1, \dots, s$ . As has been shown in the cited paper [1], it follows from relations (2) and (3) the following one:

$$(5) \quad \lim_{n \rightarrow \infty} \log \varphi_{ns} = \sum_{v=1}^j \lambda_{m_v} (e^{it_{m_v}} - 1) - \\ - 0,5 \left[ \sum_{v=j+1}^s (1 - p_{m_v}) t_{m_v}^2 - 2 \sum_{v=j+1}^{s-1} \sum_{w=v+1}^s \sqrt{p_{m_v} p_{m_w}} t_{m_v} t_{m_w} \right].$$

On the right side of (5) we have the logarithm of the characteristic function of a non-singular  $(s-1)$ -dimensional probability function of the Poisson-normal type. It is easily seen that the whole set of finite dimensional distribution functions of the Poisson-normal type given on the right side of (5) satisfies Kolmogorov's [2] consistency conditions, thus it determines a denumerably dimensional probability function of the Poisson-normal type. We have obtained the following conclusion:

If for some constants  $A_{nm}$  and  $B_{nm}$  the assumption of theorem 1a holds then it holds also for the norming of  $(x_1, x_2, x_3, \dots)$  given by (4).

<sup>3)</sup> Following a method of Cantor, we can choose a subsequence  $n_\alpha$  for which the relations (2) and (3) hold. As one can consider only this subsequence, the assumptions (2) and (3) do not restrict the generality of our considerations. The case  $\lambda_m = 0$  is — as has been shown in [1] — excluded by the assumed non-singularity of the limiting probability function.

Let us now consider the special case with equal probabilities  $p_{nm}$ , i. e. for each value of  $n$  let

$$p_{nm} = \frac{1}{r_n} \quad (m = 1, 2, \dots, r_n).$$

Then we conclude from (2) and (3) that if we use the normation (4), the limiting probability distribution will be of the Poisson type if  $r_n = O(n)$  and of the normal type if  $r_n = o(n)$ . The conclusions obtained are in accordance with intuition.

Thus if we use the  $\chi^2$ -test in a very large sample and divide the sample into classes with equal theoretical probabilities<sup>4)</sup>, the number of classes should be of order  $o(n)$ .

The theorem 1a can have the following physical interpretation. Let  $n$  denote the number of particles of gas contained in a vessel  $V$  divided into  $r_n$  parts  $V_1, V_2, \dots, V_{r_n}$ . Let us imagine a machine pressing the gas into the vessel  $V$  and at the same time dividing this vessel into more and more parts so that one can assume that  $n$  and  $r_n$  converge to infinity. Let the joint probabilities of finding  $k_m$  ( $m=1, 2, \dots, r_n$ ) particles of gas in  $V_m$  be given by (1). Then if a limiting probability distribution assumed in theorem 1a exists then 1° this limiting distribution must be of the Poisson-normal type, 2° the rapidity of the convergence of  $r_n$  to infinity is restricted by the rapidity of the convergence of  $n$  to infinity.

<sup>4)</sup> Comp. the paper of Mann and Wald [3].

#### References

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 [3] B. Mann and A. Wald, *On the choice of the number of class intervals in the application of the chi-square test*, *Annals of Mathematical Statistics* 13 (1942), p. 306-317

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