

Besides the properties (a) and (b) I assume nearly everywhere in the paper (with the exception of Theorem 1) the following third property:

$$(e) \int_0^{\infty} a(t) dt = \infty.$$

This assumption means that in the whole process there will be an infinite number of signals ( $\omega(t) \rightarrow \infty$  for  $t \rightarrow \infty$ ).

We obtain the *homogeneous* Poisson process from (a)-(e) setting  $a(t) = a > 0$ .

The purpose of this paper is to investigate the *distances between signals* in the process defined by the properties (a), (b) and (e). In section 1 I give the general definitions and theorems concerning the distributions of those distances, in section 2 I prove the limit theorem for the mean value of the fraction of distances not greater than  $y$  ( $y > 0$ ) among the initial  $n$  distances ( $n=1, 2, \dots$ ), and in section 3 I investigate the convergence of this mean value to a constant for  $n \rightarrow \infty$ .

The results obtained can be of importance in practical investigations of the non-homogeneous Poisson processes: knowing the distribution of the distances between signals it is possible to draw conclusions concerning the function  $a(t)$ . This method seems to be suitable when, for instance, the oscillation frequency of a periodic function  $a(t)$  is of the same or greater order than the mean frequency of signals.

**1. The general definitions and theorems.** I call the distance between the  $k$ -th and the  $(k+1)$ -th signals (*i. e.* the jumps of the function  $\omega(t)$ , equal to 1) the  $k$ -th *distance between signals* ( $k=1, 2, \dots$ ).

**DEFINITION 1.**  $L(y_1, \dots, y_n; \Delta y_1, \dots, \Delta y_n)$  for  $n=1, 2, \dots$ ,  $y_k \geq 0$  and  $\Delta y_k > 0$  ( $k=1, 2, \dots, n$ ) is the following event: in the whole process no less than  $n+1$  signals appear and the  $i$ -th distance between signals has a value from the interval  $(y_i, y_i + \Delta y_i)$  (for all  $i$  from 1 to  $n$ ); when we write  $\Delta y_i = \infty$ , it means that the  $i$ -th distance is greater than  $y_i$ .

**THEOREM 1.** *If the process has the properties (a) and (b), then*

$$\begin{aligned} Pr\{L(y_1, \dots, y_n; \Delta y_1, \dots, \Delta y_n)\} &= \int_0^{\infty} a(x_1) dx_1 \int_{x_1+y_1}^{x_1+y_1+\Delta y_1} a(x_2) dx_2 \times \dots \\ &\dots \times \int_{x_{n-1}+y_{n-1}}^{x_{n-1}+y_{n-1}+\Delta y_{n-1}} a(x_n) dx_n \int_{x_n+y_n}^{x_n+y_n+\Delta y_n} a(x_{n+1}) - \exp\left(\int_0^{x_{n+1}} a(t) dt\right) dx_{n+1}. \end{aligned}$$

**Proof.** In view of the properties (a) and (b) the probability that in the whole process there will be no less than  $n+1$  signals and the  $i$ -th

## On the distances between signals in the non-homogeneous Poisson stochastic process

by

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In this paper\*) I investigate a special class of the non-homogeneous Poisson stochastic processes. I denote by the random variable  $\omega(t)$  ( $t > 0$ ) the number of signals in the half-open interval  $(0, t)$ , *e. g.* the number of discharges in the Geiger-Müller counter (in this case  $t$  denotes time);  $\omega(t)$  is therefore a non-negative, non-decreasing, continuous on the right and integral valued function of  $t$ . I put  $\omega(0) = 0$ . Let us write also

$$Pr_k(t_1, t_2) = Pr\{\omega(t_2) - \omega(t_1) = k\}$$

for  $0 \leq t_1 < t_2$  and  $k=0, 1, 2, \dots$ ; it is the probability of  $k$  signals coming in the interval  $(t_1, t_2)$ .

Suppose that the process has the following properties:

(a) the random variables  $\omega(t_2) - \omega(t_1), \dots, \omega(t_{2l}) - \omega(t_{2l-1})$  are independent for  $0 \leq t_1 < t_2 \leq \dots \leq t_{2l-1} < t_{2l}$  ( $l=2, 3, \dots$ ) (process with independent increments);

$$(b) Pr_k(t_1, t_2) = \frac{\left(\int_{t_1}^{t_2} a(t) dt\right)^k}{k!} \exp\left(-\int_{t_1}^{t_2} a(t) dt\right),$$

where  $a(t)$  is a function defined for  $t \geq 0$ , non-negative and continuous.

The property (b) implies another two:

$$(c) \lim_{\Delta t \rightarrow 0} \frac{1 - Pr_0(t, t + \Delta t)}{\Delta t} = a(t);$$

$$(d) \lim_{\Delta t \rightarrow 0} \frac{1 - Pr_0(t, t + \Delta t) - Pr_1(t, t + \Delta t)}{\Delta t} = 0.$$

It is easy to see from these formulae that  $\omega(t)$  can possess only jumps equal to 1 (signals can come only singly).

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signal will appear in the interval  $(t_i, t_i + \Delta t_i)$  (for all  $i$  from 1 to  $n+1$ ), where  $0 < t_1 < t_1 + \Delta t_1 < \dots < t_{n+1} < t_{n+1} + \Delta t_{n+1}$ , is equal to

$$\begin{aligned} & Pr_0(0, t_1) Pr_1(t_1, t_1 + \Delta t_1) Pr_0(t_1 + \Delta t_1, t_2) \dots Pr_0(t_n + \Delta t_n, t_{n+1}) \times \\ & \quad \times \{1 - Pr_0(t_{n+1}, t_{n+1} + \Delta t_{n+1})\} \\ & = \left( \exp\left(-\int_0^{t_{n+1}} a(t) dt\right) - \exp\left(-\int_0^{t_{n+1} + \Delta t_{n+1}} a(t) dt\right) \right) \prod_{i=1}^n \int_{t_i}^{t_i + \Delta t_i} a(t) dt \\ & = \int_{t_1}^{t_1 + \Delta t_1} dx_1 \dots \int_{t_{n+1}}^{t_{n+1} + \Delta t_{n+1}} a(x_1) \dots a(x_{n+1}) \exp\left(-\int_0^{x_{n+1}} a(t) dt\right) dx_{n+1}. \end{aligned}$$

The function which we integrate in the last expression is therefore the conditional probability density of the appearance of the  $i$ -th signal at the point  $x_i$  (for all  $i$  from 1 to  $n+1$ ), where  $0 < x_1 < \dots < x_{n+1}$ , under the condition of the appearance in the whole process of no less than  $n+1$  signals, multiplied by the probability of this last event. Integrating this function through all the values of  $x_i$  possible in the event  $L(y_1, \dots, y_n; \Delta y_1, \dots, \Delta y_n)$ , i. e.  $x_{i-1} + y_{i-1} < x_i \leq x_{i-1} + y_{i-1} + \Delta y_{i-1}$  (for  $i$  from  $n+1$  to 2) and  $0 < x_1 < \infty$ , we get for the probability of this event the expression given in Theorem 1.

**DEFINITION 2.** The random variable  $\vartheta_k(y)$  ( $k=1, 2, \dots, y > 0$ ) is a variable equal to 1 if the  $k$ -th distance between signals is not greater than  $y$ , or equal to 0 if either the  $k$ -th distance between signals is greater than  $y$  or less than  $k+1$  signals appear in the whole process.

**THEOREM 2.** If the process has the properties (a), (b) and (e), then the mean value of the random variable  $\vartheta_k(y)$  (i. e. the probability that the  $k$ -th distance between signals will not be greater than  $y$ ) is equal to

$$E\{\vartheta_k(y)\} = 1 - \frac{1}{(k-1)!} \int_0^\infty a(x) \exp\left(-\int_0^{x+y} a(t) dt\right) \left(\int_0^x a(t) dt\right)^{k-1} dx.$$

*Proof.* Let  $y_i = 0$  (for all  $i$  from 1 to  $k-1$ ),  $y_k = y$  and  $\Delta y_i = \infty$  (for all  $i$  from 1 to  $k$ ). Then

$$\begin{aligned} E\{\vartheta_k(y)\} & = 1 - Pr\{L(y_1, \dots, y_k; \Delta y_1, \dots, \Delta y_k)\} \\ & = 1 - \int_0^\infty a(x_1) dx_1 \int_{x_1}^\infty a(x_2) dx_2 \dots \int_{x_{k-1}}^\infty a(x_k) dx_k \int_{x_k+y}^\infty a(x_{k+1}) \exp\left(-\int_0^{x_{k+1}} a(t) dt\right) dx_{k+1} \\ & = 1 - \int_0^\infty a(x_1) dx_1 \int_{x_1}^\infty a(x_2) dx_2 \dots \int_{x_{k-2}}^\infty a(x_{k-1}) dx_{k-1} \int_{x_{k-1}}^\infty a(x_k) \exp\left(-\int_0^{x_k+y} a(t) dt\right) dx_k. \end{aligned}$$

The region defined by the inequalities

$$0 < x_1 < \infty, \quad x_1 < x_2 < \infty, \quad \dots, \quad x_{k-1} < x_k < \infty$$

can also be defined by the inequalities

$$0 < x_k < \infty, \quad 0 < x_{k-1} < x_k, \quad \dots, \quad 0 < x_1 < x_2.$$

Therefore

$$\begin{aligned} E\{\vartheta_k(y)\} & = 1 - \int_0^\infty a(x_k) \exp\left(-\int_0^{x_k+y} a(t) dt\right) dx_k \int_0^{x_k} a(x_{k-1}) dx_{k-1} \dots \int_0^{x_2} a(x_1) dx_1 \\ & = 1 - \frac{1}{(k-1)!} \int_0^\infty a(x_k) \exp\left(-\int_0^{x_k+y} a(t) dt\right) \left(\int_0^{x_k} a(x_1) dx_1\right)^{k-1} dx_k, \end{aligned}$$

q. e. d.

**DEFINITION 3.** The random variable  $\Theta_n(y)$  ( $n=1, 2, \dots, y > 0$ ) is

$$\Theta_n(y) = \frac{1}{n} \sum_{k=1}^n \vartheta_k(y).$$

If no less than  $n+1$  signals appeared in the whole process (as always happens in the case of a process with property (e)), then  $\Theta_n(y)$  is the fraction of distances not greater than  $y$  among the initial  $n$  distances.

**THEOREM 3.** If the process has the properties (a), (b) and (e), then

$$E\{\Theta_n(y)\} = 1 - \frac{1}{n} \int_0^\infty a(x) \exp\left(-\int_0^{x+y} a(t) dt\right) \sum_{k=0}^{n-1} \frac{1}{k!} \left(\int_0^x a(t) dt\right)^k dx.$$

*Proof.* In view of the theorem for the mean value of the sum of random variables we obtain

$$E\{\Theta_n(y)\} = \frac{1}{n} \sum_{k=1}^n E\{\vartheta_k(y)\};$$

from this formula and Theorem 2 we deduce Theorem 3.

For the *homogeneous* process, i. e. for  $a(t) = a > 0$ , we obtain from Theorems 1, 2 and 3 respectively the known formulae:

$$(1) \quad Pr\{L(y_1, \dots, y_n; \Delta y_1, \dots, \Delta y_n)\} = \exp\left(-a \sum_{k=1}^n y_k\right) \prod_{k=1}^n (1 - e^{-a \Delta y_k}),$$

$$(2) \quad E\{\vartheta_k(y)\} = 1 - e^{-ay},$$

$$(3) \quad E\{\Theta_n(y)\} = 1 - e^{-ay}.$$

## 2. The limit theorem for $E\{\Theta_n(y)\}$ .

DEFINITION 4. For the process with properties (a), (b) and (e) I write

$$S_n(y) = 1 - \frac{1}{n} \int_0^{\tau_n} a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx,$$

where  $n=1, 2, \dots, y>0$ ;  $\tau_n$  is one of the solutions of the equation

$$\int_0^{\tau_n} a(t) dt = n.$$

If the last equation has more than one solution, then these solutions compose a closed interval in which  $a(t)=0$ ; the choice of  $\tau_n$  from this interval has therefore no significance for the value of  $S_n(y)$ .

It is easy to see from Definition 4 that  $0 < S_n(y) < 1$ .

THEOREM 4. If the process has the properties (a), (b) and (e), then

$$|E\{\Theta_n(y)\} - S_n(y)| < 1/\sqrt{2\pi n}.$$

Proof. In view of Definition 4 and Theorem 3 we obtain

$$(4) \quad S_n(y) - E\{\Theta_n(y)\} = \frac{1}{n} \int_0^{\tau_n} a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) \sum_{k=0}^{n-1} \frac{1}{k!} \left(\int_0^x a(t) dt\right)^k dx + \\ + \frac{1}{n} \int_0^{\tau_n} a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) \left\{ \exp\left(-\int_0^x a(t) dt\right) \sum_{k=0}^{n-1} \frac{1}{k!} \left(\int_0^x a(t) dt\right)^k - 1 \right\} dx.$$

For the first part of the right side of this formula we have the inequalities

$$(5) \quad 0 < \frac{1}{n} \int_0^{\tau_n} a(x) \exp\left(-\int_0^{x+y} a(t) dt\right) \sum_{k=0}^{n-1} \frac{1}{k!} \left(\int_0^x a(t) dt\right)^k dx \\ < \frac{1}{n} \int_0^{\tau_n} a(x) \exp\left(-\int_0^x a(t) dt\right) \sum_{k=0}^{n-1} \frac{1}{k!} \left(\int_0^x a(t) dt\right)^k dx \\ = \frac{1}{n} \sum_{k=0}^{n-1} \frac{1}{k!} \int_0^{\tau_n} e^{-v(x)} v(x)^k dv(x),$$

where

$$v(x) = \int_0^x a(t) dt.$$

Since, for  $k=0, 1, 2, \dots$ ,

$$\frac{1}{k!} \int_{\tau_n}^{\infty} e^{-v(x)} v(x)^k dv(x) \\ = -e^{-v(x)} \frac{v(x)^k}{k!} \Big|_{\tau_n}^{\infty} + \frac{1}{(k-1)!} \int_{\tau_n}^{\infty} e^{-v(x)} v(x)^{k-1} dv(x) = \dots = e^{-n} \sum_{l=0}^k \frac{n^l}{l!},$$

we obtain

$$(6) \quad \frac{1}{n} \sum_{k=0}^{n-1} \frac{1}{k!} \int_{\tau_n}^{\infty} e^{-v(x)} v(x)^k dv(x) = \frac{e^{-n}}{n} \sum_{k=0}^{n-1} \sum_{l=0}^k \frac{n^l}{l!} \\ = \frac{e^{-n}}{n} \sum_{l=0}^{n-1} \sum_{k=l}^{n-1} \frac{n^l}{l!} = \frac{e^{-n}}{n} \sum_{l=0}^{n-1} (n-l) \frac{n^l}{l!} \\ = \frac{e^{-n}}{n} \left( \sum_{l=0}^{n-1} \frac{n^{l+1}}{l!} - \sum_{l=0}^{n-2} \frac{n^{l+1}}{l!} \right) = \frac{n^n}{n!} e^{-n}.$$

For the second part of the right side of formula (4) we have the equalities

$$(7) \quad 0 > \frac{1}{n} \int_0^{\tau_n} a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) \times \\ \times \left\{ \exp\left(-\int_0^x a(t) dt\right) \sum_{k=0}^{n-1} \frac{1}{k!} \left(\int_0^x a(t) dt\right)^k - 1 \right\} dx \\ > \frac{1}{n} \int_0^{\tau_n} a(x) \left\{ \exp\left(-\int_0^x a(t) dt\right) \sum_{k=0}^{n-1} \frac{1}{k!} \left(\int_0^x a(t) dt\right)^k - 1 \right\} dx \\ = \frac{1}{n} \sum_{k=0}^{n-1} \frac{1}{k!} \int_0^{\tau_n} e^{-v(x)} v(x)^k dv(x) - 1.$$

Since, for  $k=0,1,2,\dots$ ,

$$\begin{aligned} & \frac{1}{k!} \int_0^{\tau_n} e^{-v(x)} v(x)^k dv(x) \\ &= -e^{-v(x)} \frac{v(x)}{k!} \Big|_0^{\tau_n} + \frac{1}{(k-1)!} \int_0^{\tau_n} e^{-v(x)} v(x)^{k-1} dv(x) = \dots = -e^{-n} \sum_{l=0}^k \frac{n^l}{l!} + 1, \end{aligned}$$

we obtain

$$(8) \quad \frac{1}{n} \sum_{k=0}^{n-1} \frac{1}{k!} \int_0^{\tau_n} e^{-v(x)} v(x)^k dv(x) - 1 = -\frac{e^{-n}}{n} \sum_{k=0}^{n-1} \sum_{l=0}^k \frac{n^l}{l!} = -\frac{n^n}{n!} e^{-n}.$$

From (4)-(8) we deduce

$$|E\{\Theta_n(y)\} - S_n(y)| < \frac{n^n}{n!} e^{-n},$$

and the inequality

$$\frac{n^n}{n!} e^{-n} < \frac{1}{\sqrt{2\pi n}}$$

completes the proof.

From Theorem 4 follows the uniform convergence of the sequence of functions  $E\{\Theta_n(y)\}$  to the sequence of functions  $S_n(y)$  in the half-axis  $y > 0$ .

Putting  $a(t) = a > 0$  in Definition 4 and comparing the obtained function  $S_n(y)$  with the function  $E\{\Theta_n(y)\}$  given by formula (3), it is easy to prove that for the homogeneous Poisson process  $S_n(y) = E\{\Theta_n(y)\}$  for all  $n$  and  $y$ .

**3. Convergence of the sequence  $E\{\Theta_n(y)\}$ .** For the homogeneous process, in view of formula (3),  $E\{\Theta_n(y)\}$  is independent of  $n$ ; it is, however, not true in the general case and the problem arises for which  $a(t)$  the sequence  $E\{\Theta_n(y)\}$  converges if  $n \rightarrow \infty$ .

Theorem 4 implies the following obvious

**COROLLARY 1.** *If the process has the properties (a), (b) and (e), then, for each  $y > 0$ , the necessary and sufficient condition for the existence of  $\lim_{n \rightarrow \infty} E\{\Theta_n(y)\}$  is that of  $\lim_{n \rightarrow \infty} S_n(y)$ ; if these limits exist, they are equal.*

This corollary enables us to replace the investigation of the convergence of the sequence  $E\{\Theta_n(y)\}$  by such of the convergence of the sequence  $S_n(y)$ , which we shall now consider. Since we have defined  $S_n(y)$  only for the processes with properties (a), (b) and (e) we shall not assume *explicitly* in the following that the process has those properties.

**LEMMA.** *The necessary and sufficient condition for the existence, for some  $y > 0$ , of the limit  $\lim_{n \rightarrow \infty} S_n(y)$  is that of the limit*

$$\lim_{z \rightarrow \infty} \left( 1 - \frac{\int_0^z a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx}{\int_0^z a(t) dt} \right);$$

if these limits exist, they are equal.

Proof. Let<sup>1)</sup>

$$n(z) = \left[ \int_0^z a(t) dt \right];$$

for  $z \geq \tau_1$ , i. e.  $n(z) \geq 1$ , we have

$$(9) \quad \frac{\int_0^z a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx}{\int_0^z a(t) dt} = \frac{n(z) \int_0^{\tau_{n(z)}} a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx}{\int_0^z a(t) dt} + \frac{\int_{\tau_{n(z)}}^z a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx}{\int_0^z a(t) dt}.$$

If  $z \rightarrow \infty$ , then  $n(z) \rightarrow \infty$  successively throughout all natural numbers; since

$$\lim_{z \rightarrow \infty} \frac{n(z)}{\int_0^z a(t) dt} = 1, \quad \lim_{z \rightarrow \infty} \frac{\int_{\tau_{n(z)}}^z a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx}{\int_0^z a(t) dt} = 0,$$

we deduce our Lemma from formula (9).

**THEOREM 5.** *If, for some  $y > 0$  and  $C' \geq 0$ , the limit*

$$\lim_{n \rightarrow \infty} S_n(y)$$

*exists in the process defined by the function  $a(t+C')$ , then this limit exists and is the same in the process defined by the function  $a(t+C)$ , for each  $C \geq 0$ .*

<sup>1)</sup> By the symbol  $[x]$  I denote, in this proof and in the proof of Theorem 8, the function "entier  $x$ ".

Proof. In virtue of the Lemma and by transformations of integrals we get

$$\begin{aligned} \lim_{n \rightarrow \infty} S_n(y)_{a(t+C)} &= 1 - \lim_{z \rightarrow \infty} \frac{\int_0^z a(x+C') \exp\left(-\int_x^{x+y} a(t+C') dt\right) dx}{\int_0^z a(t+C') dt} \\ &= 1 - \lim_{z \rightarrow \infty} \frac{\int_{C'-C}^{z+C'-C} a(x+C) \exp\left(-\int_x^{x+y} a(t+C) dt\right) dx}{\int_{C'-C}^{z+C'-C} a(t+C) dt} \\ &= 1 - \lim_{z \rightarrow \infty} \frac{\int_0^{z+C'-C} a(x+C) \exp\left(-\int_x^{x+y} a(t+C) dt\right) dx}{\int_0^{z+C'-C} a(t+C) dt} \\ &= \lim_{n \rightarrow \infty} S_n(y)_{a(t+C)}, \end{aligned}$$

q. e. d.

**THEOREM 6.** *If, for some  $y > 0$ ,  $\lim_{n \rightarrow \infty} S_n(y)$  exists in the process defined by the function  $a(t)$ , then this limit exists and is the same in the process defined by the function  $a(t) + b(t)$ , where  $b(t)$  is any continuous function, defined for  $t \geq 0$ , such that  $a(t) + b(t) \geq 0$  for all  $t \geq 0$  and the integral*

$$\int_0^\infty |b(t)| dt$$

is finite.

Proof. Denote by  $M_C$  and  $m_C$  respectively the upper and the lower limits of the function

$$f(x) = \int_x^{x+y} b(t) dt$$

for  $x \geq C$ , where  $C \geq 0$ ; for any  $z \geq C \geq \tau_1$  we have

$$\begin{aligned} e^{-M_C} \frac{\int_0^z a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx}{\int_0^z a(t) dt} &< \frac{\int_0^z a(x) \exp\left(-\int_x^{x+y} \{a(t) + b(t)\} dt\right) dx}{\int_0^z a(t) dt} \\ &\leq e^{-m_C} \frac{\int_0^z a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx}{\int_0^z a(t) dt} + \frac{\int_0^z a(x) \exp\left(-\int_x^{x+y} \{a(t) + b(t)\} dt\right) dx}{\int_0^z a(t) dt} \end{aligned}$$

If  $z \rightarrow \infty$ , the first and the last expressions converge in virtue of the Lemma to

$$e^{-M_C} \{1 - \lim_{n \rightarrow \infty} S_n(y)_{a(t)}\} \quad \text{and} \quad e^{-m_C} \{1 - \lim_{n \rightarrow \infty} S_n(y)_{a(t)}\}$$

respectively. Since  $C$  is arbitrary and, as follows from the existence of the integral  $\int_0^\infty |b(t)| dt$ ,  $\lim_{C \rightarrow \infty} M_C = \lim_{C \rightarrow \infty} m_C = 0$ , we obtain

$$(10) \quad 1 - \lim_{n \rightarrow \infty} S_n(y)_{a(t)} = \lim_{z \rightarrow \infty} \frac{\int_0^z a(x) \exp\left(-\int_x^{x+y} \{a(t) + b(t)\} dt\right) dx}{\int_0^z a(t) dt}.$$

We also have

$$(11) \quad \left| \int_0^z b(t) dt \right| \leq \int_0^z |b(t)| dt \leq \int_0^\infty |b(t)| dt$$

and

$$(12) \quad \left| \int_0^z b(x) \exp\left(-\int_x^{x+y} \{a(t) + b(t)\} dt\right) dx \right| \leq \int_0^z |b(x)| \exp\left(-\int_x^{x+y} \{a(t) + b(t)\} dt\right) dx \leq \int_0^\infty |b(x)| dx.$$

From (10), (11), (12) and the Lemma we deduce

$$1 - \lim_{n \rightarrow \infty} S_n(y)_{a(t)}$$

$$= \lim_{z \rightarrow \infty} \frac{\int_0^z a(x) \exp\left(-\int_x^{x+y} \{a(t) + b(t)\} dt\right) dx + \int_0^z b(x) \exp\left(-\int_x^{x+y} \{a(t) + b(t)\} dt\right) dx}{\int_0^z a(t) dt + \int_0^z b(t) dt}$$

$$= 1 - \lim_{n \rightarrow \infty} S_n(y)_{a(t)+b(t)}, \quad \text{q. e. d.}$$

**THEOREM 7.** *If, for some  $y > 0$ , the limit*

$$\lim_{n \rightarrow \infty} S_n(y)$$

exists in the process defined by the function  $a(t)$ , then also, for each  $C > 0$ , the limit

$$\lim_{n \rightarrow \infty} S_n(y/C),$$

equal to the first, exists in the process defined by the function  $Ca(Ct)$ .

Proof. In virtue of the Lemma and by transformations of integrals we obtain

$$\begin{aligned} \lim_{n \rightarrow \infty} S_n(y)_{a(t)} &= 1 - \lim_{z \rightarrow \infty} \frac{\int_0^z a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx}{\int_0^z a(t) dt} \\ &= 1 - \lim_{z \rightarrow \infty} \frac{\int_0^{z/C} C a(Cx) \exp\left(-\int_x^{x+y/C} C a(Ct) dt\right) dx}{\int_0^{z/C} C a(Ct) dt} = \lim_{n \rightarrow \infty} S_n(y/C)_{Ca(Ct)}, \end{aligned}$$

q. e. d.

THEOREM 8. If, for some  $y > 0$ , the limit

$$\lim_{z \rightarrow \infty} \int_z^{z+y} a(t) dt = A_y$$

exists (in particular we may have  $A_y = \infty$ ), then the function  $a(t)$  has the mean value

$$a = \lim_{z \rightarrow \infty} \frac{1}{z} \int_0^z a(t) dt = \frac{A_y}{y},$$

and the limit

$$\lim_{n \rightarrow \infty} S_n(y) = 1 - e^{-ay}$$

exists.

Proof. For  $z \geq y$  we have

$$\frac{\sum_{k=1}^{[z/y]} \int_{(k-1)y}^{ky} a(t) dt}{([z/y]+1)y} \leq \frac{1}{z} \int_0^z a(t) dt \leq \frac{\sum_{k=1}^{[z/y]+1} \int_{(k-1)y}^{ky} a(t) dt}{[z/y]y}.$$

Because of the convergence of the sequence

$$C_k = \int_{(k-1)y}^{ky} a(t) dt \quad (k=1, 2, \dots)$$

to  $A_y$ , the sequence of the arithmetical means of the sequence  $C_k$  also converges to this limit; therefore, if  $z \rightarrow \infty$ , we deduce from the inequalities obtained that the limit

$$\lim_{z \rightarrow \infty} \frac{1}{z} \int_0^z a(t) dt = \frac{A_y}{y}$$

exists.

Let us consider the function  $g(t) = a(t) + e^{-t} > 0$ . By virtue of de l'Hospital's rule and the Lemma we get

$$\begin{aligned} e^{-Ay} &= \lim_{z \rightarrow \infty} \exp\left(-\int_z^{z+y} a(t) dt + e^{-z}(e^{-y} - 1)\right) = \lim_{z \rightarrow \infty} \exp\left(-\int_z^{z+y} g(t) dt\right) \\ &= \lim_{z \rightarrow \infty} \frac{g(z) \exp\left(-\int_z^{z+y} g(t) dt\right)}{g(z)} = \lim_{z \rightarrow \infty} \frac{\int_0^z g(x) \exp\left(-\int_x^{x+y} g(t) dt\right) dx}{\int_0^z g(t) dt} \\ &= 1 - \lim_{n \rightarrow \infty} S_n(y)_{g(t)}, \end{aligned}$$

and since  $\int_0^\infty e^{-t} dt = 1$ , we have from Theorem 6

$$e^{-Ay} = 1 - \lim_{n \rightarrow \infty} S_n(y)_{a(t)},$$

which completes the proof.

Theorem 8 implies the following obvious

COROLLARY 2. If the limit  $\lim_{t \rightarrow \infty} a(t) = a$  exists (in particular we may have  $a = \infty$ ), then  $a$  is the mean value of function  $a(t)$  and for each  $y > 0$  the limit

$$\lim_{n \rightarrow \infty} S_n(y) = 1 - e^{-ay}$$

exists.

THEOREM 9. If the function  $a(t)$  is periodic with period  $T$ , then for each  $y > 0$  the limit

$$\lim_{n \rightarrow \infty} S_n(y) = 1 - \frac{\int_0^T a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx}{\int_0^T a(t) dt}$$

exists.

Proof. In view of the periodicity of the functions which we integrate, for every  $y > 0$  the limits

$$\begin{aligned} \lim_{z \rightarrow \infty} \frac{1}{z} \int_0^z a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx &= \frac{1}{T} \int_0^T a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx, \\ \lim_{z \rightarrow \infty} \frac{1}{z} \int_0^z a(t) dt &= \frac{1}{T} \int_0^T a(t) dt > 0, \end{aligned}$$

exist. Therefore we obtain by virtue of the Lemma

$$1 - \frac{\int_0^T a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx}{\int_0^T a(t) dt} = 1 - \lim_{z \rightarrow \infty} \frac{z^{-1} \int_0^z a(x) \exp\left(-\int_x^{x+y} a(t) dt\right) dx}{z^{-1} \int_0^z a(t) dt} \\ = \lim_{n \rightarrow \infty} S_n(y), \text{ q. e. d.}$$

S. Hartman has proved (by a somewhat different method from the one employed here) the following

**THEOREM 10.** *If the function  $a(t)$  is, for  $t \geq 0$ , equal to a certain uniformly almost periodic function, then for every  $y > 0$  the limit  $\lim_{n \rightarrow \infty} S_n(y)$  exists.*

*Proof.* Denote by  $a_0(t)$  the uniformly almost periodic function which is equal to  $a(t)$  for  $t \geq 0$ . Because of its being uniformly almost periodic the function  $a_0(t)$  is bounded. Hence, there exists a number  $M > 0$ , such that  $-M \leq a_0(t) \leq M$  for all  $t$ . We have also

$$\int_x^{x+y} a_0(t) dt = \int_0^y a_0(t) dt + \int_0^x \{a_0(t+y) - a_0(t)\} dt.$$

Since

$$-My \leq \int_x^{x+y} a_0(t) dt \leq My,$$

the integral  $\int_0^x \{a_0(t+y) - a_0(t)\} dt$  is the bounded indefinite integral of a uniformly almost periodic function, *i. e.* it is a uniformly almost periodic function of  $x$ . Therefore  $\int_x^{x+y} a_0(t) dt$  is also a uniformly almost periodic function of  $x$ .

Since

$$\exp\left(-\int_x^{x+y} a_0(t) dt\right) = \sum_{k=0}^{\infty} \frac{(-1)^k}{k!} \left(\int_x^{x+y} a_0(t) dt\right)^k,$$

with the series uniformly convergent for all  $x$  in view of the boundedness of  $\int_x^{x+y} a_0(t) dt$ ,  $\exp\left(-\int_x^{x+y} a_0(t) dt\right)$  is a uniformly almost periodic function of  $x$ .

In virtue of the mean value theorem for uniformly almost periodic functions, the limits

$$\lim_{z \rightarrow \infty} \frac{1}{z} \int_0^z a_0(x) \exp\left(-\int_x^{x+y} a_0(t) dt\right) dx$$

and

$$\lim_{z \rightarrow \infty} \frac{1}{z} \int_0^z a_0(t) dt$$

exist; moreover the last limit is greater than zero, because for each  $\varepsilon > 0$  there exists  $l > \tau_1$ , such that

$$\int_{(k-1)l}^{kl} a_0(t) dt > 1 - \varepsilon \tau_1$$

for all natural  $k$ . Therefore the limit

$$\lim_{z \rightarrow \infty} \frac{z^{-1} \int_0^z a_0(x) \exp\left(-\int_x^{x+y} a_0(t) dt\right) dx}{z^{-1} \int_0^z a_0(t) dt}$$

exists, and in view of the Lemma we deduce the existence of the limit  $\lim_{n \rightarrow \infty} S_n(y)$ , *q. e. d.*

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