

## The distribution of divisors of $N!$

by

MICHAEL D. VOSE (Austin, Tex.)

**Introduction.** The divisors of  $j!$  have an asymptotic distribution in which the set of prime numbers is embedded. An explicit formula for this distribution is given which leads to a sequence of splines converging to its density. Approximation, rate of convergence, and large deviations are also considered.

**Results.** Let  $X_j$  be a random variable uniformly distributed over the set  $\{\log d: d|j!\}$  and let  $F_j$  be the normalized (to have expectation zero and variance one) distribution function for  $X_j$ . Let  $p_i$  be the  $i$ th prime and let  $\chi_I$  be the indicator of the interval  $I$ .

**THEOREM 1.** *The sequence  $F_j$  converges completely to the distribution  $\psi$  having density  $q$  represented by the infinite convolution*

$$q = X_1 * X_2 * \dots,$$

where

$$X_i = (2\xi_i)^{-1} \chi_{[-\xi_i, \xi_i]},$$

$$\xi_i = \frac{\sigma \log p_i}{p_i - 1},$$

$$\sigma = \left( \frac{1}{3} \sum_p \left( \frac{\log p}{p-1} \right)^2 \right)^{-1/2}.$$

Moreover,  $\sup_x |F_j(x) - \psi(x)| \ll_\varepsilon j^{-1/3+\varepsilon}$  for any  $\varepsilon > 0$ .

Let  $q_N = X_1 * \dots * X_N$ . The sequence of splines  $\{q_N\}_1^\infty$  converges to  $q$  and is given explicitly by

**THEOREM 2.** For  $N > 1$ ,

$$q_N = A_N \sum_{j=1}^{2^{N-1}} s_N(j) (\langle \theta_N(j) + x \rangle^{N-1} + \langle \theta_N(j) - x \rangle^{N-1}),$$

where

$$A_N = \frac{1}{4(N-1)!} \prod_{j=1}^N \xi_j^{-1},$$

$$s_N(j) = 2^{1-N} (-1)^{\sum_{0 < i < N} \varepsilon_i(j)},$$

$$\theta_N(j) = \xi_j + \sum_{0 < i < N} (-1)^{\varepsilon_i(j)} \xi_{i+1},$$

$$\varepsilon_i(j) = [2 \{j2^{-i}\}],$$

$$\langle x \rangle^M = x^M \operatorname{sgn}(x).$$

Since logarithms of primes are linearly independent over  $\mathcal{Q}$ , the nodes of the spline  $\varrho_N$  consist of the  $2^N$  distinct points  $\{\pm \theta_N(j)\}_1^{2^N-1}$ . Moreover,  $\varrho_N$  is supported on an interval symmetric about the origin of length  $\ll \log N$ . Since the complexity of  $\varrho_N$  makes it difficult to calculate, the representations

$$\hat{\psi}(t) = \prod_{k>0} \frac{\sin(\xi_k t)}{\xi_k t} = \exp \left\{ - \sum_{n>0} \frac{4^n B_n}{(2n)(2n)!} \sum_{k>0} (\xi_k t)^{2n} \right\}$$

(see [2]) for the characteristic function of  $\psi$  were used instead to obtain the following result which is stated as a conjecture (this computation has not been independently verified).

CONJECTURE. Let  $f(x) = \exp(-.954 - .434x^2 - .011x^4)$ . Then

$$|f(x) - \varrho(x)| < .001 \quad \text{for } |x| < 3.$$

The probability distribution  $\psi$  concentrates mass about the origin more so than does the Gaussian distribution. We have the following result as  $x \rightarrow \infty$ :

THEOREM 3. Both  $1 - \psi(x)$  and  $\varrho(x)$  are  $o(\exp(-e^{(\sigma^{-1}-\varepsilon)x}))$  for any  $\varepsilon > 0$ .

#### Demonstrations.

Proof of Theorem 1. Define  $\mu_n(j)$  for positive integer  $n$  by

$$\mu_n = (12^{-1} \sum_{p^a \parallel j} ((\alpha+1)^n - 1)(\log p)^n)^{1/n}.$$

The convergence of  $F_j$  to  $\psi$  follows from Theorem 1 of [2] which in the present case reduces to:

THEOREM. A necessary and sufficient condition for  $F_j$  to converge to a distribution  $\psi$  is that for each  $n$  the limits  $a_n = \lim_{j \rightarrow \infty} \mu_{2n}/\mu_2$  exist. In this case  $\hat{\psi}$  is entire and is represented in the disk  $|z| < 1/4$  by

$$\hat{\psi}(z) = \exp \left( - \sum_{n>0} \frac{6B_n}{n(2n)!} (a_n z)^{2n} \right),$$

where

$$B_n = 4n \int_0^{\infty} (e^{2\pi t} - 1)^{-1} t^{2n-1} dt$$

are the Bernoulli numbers.

The limits  $a_n$  are computed from the prime factorization

$$j! = p_1^{\alpha_1} \dots p_k^{\alpha_k}, \quad \text{where } \alpha_i = \sum_{t>0} [j p_i^{-t}].$$

It follows that

$$(1.1) \quad \alpha_i = \frac{j}{p_i - 1} + O\left(\frac{\log j}{\log p_i}\right),$$

and, for any  $\varepsilon > 0$ ,

$$(1.2) \quad (\mu_{2n}(j))^{2n} = \frac{j^{2n}}{12} \sum_p \left(\frac{\log p}{p-1}\right)^{2n} (1 + O_\varepsilon(j^{\varepsilon-1}))^{2n}.$$

Hence the limit distribution  $\psi$  exists, and  $\hat{\psi}$  is represented near the origin by

$$\begin{aligned} \hat{\psi}(t) &= \exp \left\{ - \sum_{n>0} \frac{4^n B_n}{(2n)(2n)!} \sum_{k>0} (\xi_k t)^{2n} \right\} \\ &= \exp \left\{ \sum_{k>0} \int_0^{\xi_k |t|} \cot x - x^{-1} dx \right\} \\ &= \prod_{k>0} \frac{\sin(\xi_k t)}{\xi_k t}. \end{aligned}$$

Therefore,  $\varrho = X_1 * X_2 * \dots$

To estimate the rate of convergence of  $F_j$  to  $\psi$ , we use the Berry-Esseen inequality [1]:

For all  $T > 0$ ,

$$(1.3) \quad \sup_x |F_j(x) - \psi(x)| \ll \frac{1}{T} + \int_{-T}^T \frac{|\hat{F}_j(t) - \hat{\psi}(t)|}{|t|} dt.$$

We will use the following representations of  $\hat{F}_j$ :

$$\hat{F}_j(t) = \exp \left\{ - \sum_{n=1}^{\infty} \frac{6B_n}{n(2n)!} (\mu_{2n}/\mu_2)^{2n} t^{2n} \right\} = \prod_{i=1}^k R_i(t),$$

where

$$R_i(t) = \frac{\sin\left(\frac{t}{2\mu_2}(\alpha_i + 1)\log p_i\right)}{(\alpha_i + 1)\sin\left(\frac{t}{2\mu_2}\log p_i\right)},$$

see [2]. From (1.1) and (1.2) follows the estimate

$$(1.4) \quad R_i(t) = \frac{\sin(\xi_i t)}{\xi_i t} + O_\varepsilon\left(\frac{p_i}{j^{1-\varepsilon}\log p_i} + \frac{t^2 \log^2 p_1}{j^{2-\varepsilon}}\right).$$

Let  $0 < \eta < 1/2$  and suppose  $1 \ll t \leq j^\eta$ . Since  $\xi_i \asymp t^{-1}$ , Sterling's formula applied to the product representations of  $\hat{F}_j$  and  $\hat{\psi}$  gives

$$|\hat{F}_j(t)| + |\hat{\psi}(t)| \ll \sqrt{t}e^{-t}.$$

Assuming the further condition that  $t^2 = O(j^\eta)$ , and using

$$(a \sin x)^{-1} \sin ax = 1 + O((1+a^2)x^2),$$

we have

$$\prod_{i=j^\eta}^k R_i(t) = 1 + O_\varepsilon(T^2 j^{-\eta}).$$

Since the tail of the product representing  $\hat{\psi}$  is similarly small, (1.4) applied to the product representation for  $\hat{F}_j$  allows us to conclude:

$$\begin{aligned} |\hat{F}_j(t) - \hat{\psi}(t)| &\ll \sqrt{t}e^{-t} && \text{for } 1 \ll t \leq j^\eta, \\ |\hat{F}_j(t) - \hat{\psi}(t)| &\ll_\varepsilon \frac{t^{5/2}e^{-t}}{j^\eta} + j^{\varepsilon+2\eta-1} && \text{for } 1 \ll t^2 = O(j^\eta). \end{aligned}$$

We base our estimate of  $|\hat{F}_j - \hat{\psi}|$  for small values of  $t$  on (1.4). Using (1.2) we have

$$(\mu_{2n}/\mu_2)^{2n} = \frac{4^n}{12} \sum_{k>0} (\xi_k (1 + O_\varepsilon(j^{\varepsilon-1})))^{2n},$$

hence

$$\hat{F}_j(t) = \hat{\psi}(t(1 + O(j^{\varepsilon-1}))).$$

Since  $\hat{\psi}'(t)$  is Lipschitz for small  $t$ , and since  $\hat{\psi}'(0) = 0$ , we obtain

$$|\hat{F}_j(t) - \hat{\psi}(t)| \ll t^2 j^{\varepsilon-1} \quad \text{for } |t| < \frac{1}{4}.$$

Using our estimates of  $|\hat{F}_j - \hat{\psi}|$  in (1.3) with  $T = j^\eta$  and  $\eta = \frac{1}{3} + \varepsilon$  completes the proof. ■

Proof of Theorem 2. Let  $t_N(x) = \left(-\frac{d}{dx}\right)^N \cos x$ . Since

$$\prod_{j=1}^N \sin \xi_j = \sum_{j=1}^{2^{N-1}} s_N(j) t_N(\theta_N(j)),$$

Fourier inversion gives

$$(2.1) \quad \varrho_N(x) = \lim_{\varepsilon \rightarrow 0} \frac{4(N-1)! A_N}{\pi} \sum_{j=1}^{2^{N-1}} s_N(j) \int_{\varepsilon}^{\infty} \cos(ux) t_N(\theta_N(j)u) \frac{du}{u^N}.$$

Assume  $N > 1$  is odd, say  $N = 2n+1$ . Let  $\eta^+ = |\theta_N(j) + x|$  and  $\eta^- = |\theta_N(j) - x|$ . Then the integral in (2.1) is

$$(2.2) \quad \left( \langle \theta_N(j) + x \rangle^{N-1} \int_{\eta^+ \varepsilon}^{\infty} \frac{\sin u}{2u^{2n+1}} du + \langle \theta_N(j) - x \rangle^{N-1} \int_{\eta^- \varepsilon}^{\infty} \frac{\sin u}{2u^{2n+1}} du \right) (-1)^n.$$

Since

$$\int \frac{\sin u}{u^{2n+1}} du = \left( H_n(u) + \frac{1}{(2n)!} \int \frac{\sin u}{u} du \right) (-1)^n$$

where

$$H_n(u) = \frac{1}{u(2n)!} \sum_{k=0}^{n-1} \frac{(-1)^k (2k)!}{u^{2k}} \left( \cos u + \frac{2k+1}{u} \sin u \right),$$

(2.2) becomes

$$\begin{aligned} \langle \theta_N(j) + x \rangle^{N-1} &\left( \frac{\pi}{4(2n)!} + o(1) - H_n(\eta^+ \varepsilon) \right) \\ &+ \langle \theta_N(j) - x \rangle^{N-1} \left( \frac{\pi}{4(2n)!} + o(1) - H_n(\eta^- \varepsilon) \right). \end{aligned}$$

Therefore  $\varrho_N$  is represented by

$$(2.3) \quad A_N \sum_{j=1}^{2^{N-1}} s_N(j) (\langle \theta_N(j) + x \rangle^{N-1} + \langle \theta_N(j) - x \rangle^{N-1}) - \lim_{\varepsilon \rightarrow 0} \frac{4(N-1)! A_N}{\pi} \sum_{j=1}^{2^{N-1}} s_N(j) (\langle \theta_N(j) + x \rangle^{N-1} H_n(\eta^+ \varepsilon) + \langle \theta_N(j) - x \rangle^{N-1} H_n(\eta^- \varepsilon)).$$

Now view the  $\xi_i$  as indeterminates so that  $\varrho_N$  is a function of the  $\xi_i$  and of  $x$ . Notice that if  $\xi_i$  and  $x$  are algebraic, then both  $\varrho_N(x)$  and the first term of



(2.3) are algebraic. Since  $\pi$  is transcendental, it follows that the limit as  $\varepsilon \rightarrow 0$  of the sum in the second term of (2.3) is either transcendental or zero. Since  $H_n(u)$  is meromorphic, with rational coefficients in its Laurent expansion about zero, this limit must therefore be zero. ■

Before proving Theorem 3, we establish the following

LEMMA. For any  $\varepsilon > 0$ ,  $\hat{q}(iy) \ll_\varepsilon \exp((\sigma + \varepsilon)y \log y)$  as  $y \rightarrow \infty$ .

Proof of the Lemma. We have

$$\hat{q}(iy) \ll \prod_{k=1}^{\infty} (\xi_k y)^{-1} \sinh(\xi_k y) = \exp\left(\sum_{k=1}^{\infty} \log((\xi_k y)^{-1} \sinh(\xi_k y))\right).$$

Since  $x^{-1} \sinh x < e^x$  for  $x \geq \frac{1}{2}$ , and since  $\xi_k \sim \sigma k^{-1}$ , we have

$$\sum_{k \leq 3\sigma y/2} \log((\xi_k y)^{-1} \sinh(\xi_k y)) \leq (\sigma + \varepsilon)y \log y,$$

for any  $\varepsilon > 0$  provided  $y$  is sufficiently large (depending on  $\varepsilon$ ). If  $|x| < 1$ , the  $\log\left(\frac{\sinh x}{x}\right) \ll x^2$ , so that

$$\sum_{k > 3\sigma y/2} ((\xi_k y)^{-1} \sinh(\xi_k y)) \ll y.$$

Combining these estimates completes the proof of the lemma. ■

Proof of Theorem 3. Let  $f(x) = \chi_{[y, \infty)}(x) e^{-\alpha(y)x}$ ,  $g(x) = e^{\alpha(y)x} \varrho(x)$ ,  $h_\lambda(x) = \frac{1}{2\lambda} \chi_{[-\lambda, \lambda]}(x)$ , where  $\alpha(y)$  is a function to be chosen later, and  $\lambda$  is positive parameter. It follows that

$$1 - \psi(y) = \int f * h_\lambda(x) g(x) dx + \int g(x)(f(x) - f * h_\lambda(x)) dx,$$

which by the Parseval identity applied to the first integral is

$$\int \left(\frac{1}{\sqrt{2\pi}} \frac{e^{-\alpha(y)y + ity}}{\alpha(y) - it}\right) \left(\frac{1}{\sqrt{2\pi}} \frac{\sin(t\lambda)}{t\lambda}\right) \hat{q}(t + i\alpha(y)) dt + \int e^{\alpha(y)x} \varrho(x) H(x) dx$$

where

$$H(x) = \frac{1}{2\lambda} \int_{-\lambda}^{\lambda} (f(x) - f(x-t)) dt.$$

Note that

$$H(x) \begin{cases} = 0 & \text{if } x < y - \lambda, \\ = e^{-\alpha(y)x} \left(1 - \frac{\sinh(\lambda\alpha(y))}{\lambda\alpha(y)}\right) & \text{if } x > y + \lambda, \\ \ll e^{-\alpha(y)(y-\lambda)} & \text{if } y - \lambda \leq x \leq y + \lambda. \end{cases}$$

Assuming that  $\alpha(y) \rightarrow \infty$  as  $y \rightarrow \infty$ ,  $\lambda = o(\alpha(y)^{-1})$ , and estimating  $\hat{q}(i\alpha(y))$  by the lemma produces

$$1 - \psi(y) \ll_\varepsilon \lambda^{-1} \exp(-\alpha(y)(y - (\sigma + \varepsilon) \log \alpha(y))) + \lambda \varrho(y - \lambda).$$

Choosing  $\alpha(y) = \exp((\sigma + \varepsilon)^{-1} y - 1)$ , and noting that  $\varrho(y + 1) < 1 - \psi(y)$ , we obtain

$$\varrho(y + 1) \ll_{\varepsilon_0} \lambda^{-1} \exp(-\exp((\sigma^{-1} - \varepsilon_0)y)) + \lambda \varrho(y - \lambda),$$

for any  $\varepsilon_0 > 0$ . Choosing

$$\lambda = \exp(-\frac{1}{2} \exp((\sigma^{-1} - \varepsilon_0)y))$$

completes the proof since  $\varepsilon_0 > 0$  was arbitrary. ■

References

- [1] W. Feller, *An Introduction to Probability Theory and its Applications*, vol. II, Wiley, New York 1966.
- [2] M. Vose, *Limit theorems for divisor distributions*, Proc. Amer. Math. Soc. 95, Number 4, (1985), pp. 505-511.

Received on 9.5.1986

(1634)